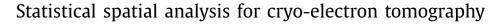
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ABSTRACT

Cryo-electron tomography (cryo-ET) is uniquely suited to precisely localize macromolecular complexes *in situ*, that is in a close-to-native state within their cellular compartments, in three-dimensions at high resolution. Point pattern analysis (PPA) allows quantitative characterization of the spatial organization of particles. However, current implementations of PPA functions are not suitable for applications to cryo-ET data because they do not consider the real, typically irregular 3D shape of cellular compartments and molecular complexes. Here, we designed and implemented first and the second-order, uni- and bivariate PPA functions in a Python package for statistical spatial analysis of particles located in three dimensional regions of arbitrary shape, such as those encountered in cellular cryo-ET imaging (PyOrg).

To validate the implemented functions, we applied them to specially designed synthetic datasets. This allowed us to find the algorithmic solutions that provide the best accuracy and computational performance, and to evaluate the precision of the implemented functions. Applications to experimental data showed that despite the higher computational demand, the use of the second-order functions is advantageous to the first-order ones, because they allow characterization of the particle organization and statistical inference over a range of distance scales, as well as the comparative analysis between experimental groups comprising multiple tomograms.

Altogether, PyOrg is a versatile, precise, and efficient open-source software for reliable quantitative characterization of macromolecular organization within cellular compartments imaged *in situ* by cryo-ET, as well as to other 3D imaging systems where real-size particles are located within regions possessing complex geometry.

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1. Introduction

The last decades of research in cell biology have revealed that cellular processes are performed by groups of interacting macromolecules in a crowded environment. This is in contrast to earlier models where macromolecules were considered to exist as isolated objects floating randomly in the cytoplasm. Therefore the analysis of their organization within their native cellular compartments can provide quantitative information that can be used to describe the mechanisms underlying macromolecular interactions. This information has paramount importance to gain a deeper understanding of various cellular interactions [1,2].

Cryo-Electron Tomography (cryo-ET) is a unique imaging technique capable of producing 3D views of large portions of cells at a resolution that is sufficiently high to localize and identify macromolecular complexes [3]. In cryo-ET, biological samples are vitrified to preserve their natural molecular organization and are imaged by electron microscopy in the fully-hydrated vitrified state, thus enabling the study of cells in a close-to-native state at high resolution [4,5].

Point patterns analysis (PPA) is a branch of statistics devoted to describing quantitatively point patterns in space. PPA has been used extensively in experimental fields such as ecology [6], so-cial sciences [7,8], and more recently in biology [9,10]. Among PPA

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methods, first-order functions describe clustering of points and typically determine a characteristic distance scale, whereas secondorder functions determine spatial correlations between points at a range of distances. Ripley's functions K (or its linearization, L) and O [6] are the most used tools for second-order PPA. They have been used for analyzing experimental biological data obtained from light and electron microscopy [10-12]. Current implementations of the PPA methods were either developed for 2D spaces [13], or their application in 3D is limited to sphere-shaped structures within cubic volumes [11,14]. However, these approaches have two limitations, rendering their application impractical for cellular cryo-ET data: (i) cellular compartments have complex geometry so, there are no analytical solutions for border compensation, unlike for the most simple geometries [14,15], and (ii) proteins have specific shapes that can not be properly represented by points or spheres. Recently Voronoi tessellation has been used to describe the particle organization in super-resolution optical microscopy [16], but contrarily to Ripley's functions it does not allow comparison between datasets having different density of points.

The authors have already applied PPA functions to successfully solve some long-standing biological questions. Firstly, in [2], we applied a modification of Ripley's function *O* to prove that Rubisco complex has a liquid-like organization within the pyrenoid organelle, thus discarding crystalline models proposed previously. This finding was also remarkable as cryo-ET data was used to successfully analyze a phase-separated compartment *in situ* at nanometric resolution for the first time. We also showed that the Arp2/3 complex within actin waves modifies its clustered organization depending on the wave-phase [17]. Recently, we demonstrated that Ripley's functions can be used to characterize nano-domains formed by synaptic membrane-bound complexes [18].

Here we present an implementation of the first and second order PPA functions where both particles and cellular compartments can be of arbitrary shape. Consequently, these functions are applicable to cryo-ET images of cellular samples, that is they are suitable for the spatial distribution analysis of individual proteins or macromolecular complexes localized in any kind of cellular compartment (cytoplasm, organelle lumen, membrane, etc). These numerical calculations are required to precisely calculate the computationally intensive second order PPA functions. We also present a parallel implementation suitable for processing cryo-ET datasets that takes advantage of modern multiprocessor architectures [19]. In addition to univariate PPA functions, we also implemented the bivariate versions of PPA functions, which enable the colocalization analysis between different proteins and macromolecules.

We present several applications to synthetic datasets in order to compare different numerical methods and find the most appropriate ones. Finally, we also applied the methods we implemented on real Cryo-ET datasets to validate their real-case usability and to justify the necessity for an accurate implementation of secondorder PPA function.

2. Approach

In this section, we show the implementation details relevant for the first and second order PPA functions, for univariate and bivariate cases.

2.1. Design

2.1.1. Monovariate first-order analysis

Nearest neighbor function *G*, spatial contact distribution function *F*, and their combination *J* belong to the first order PPA functions. Location of proteins or macromolecular complexes of interest (particles) in a tomogram (3D image) is defined by their spatial coordinates $\mathbf{X} = \{\mathbf{x}_i \in \mathcal{V} \forall i = (1, ..., n)\}$, where *n* is the number of

particles and $\mathcal{V} \in \mathbb{R}^d$ is the Volume of Interest (VOI), that is a subspace of the tomogram where the particles are located. A VOI typically represents an organelle or a distinct cellular compartment.

Function *G* is defined as:

$$G(r) = \int_0^r p_G(r) dr \tag{1}$$

where $p_G(r)$ is the probability distribution function of nearest neighbor distances r of all particles in **X**.

Function *F* is similarly defined, except that it requires $p_F(r)$, the probability distribution function of nearest neighbor distances among points in sets **X** and **X^{CSR}** $\in \mathcal{V}$, where **X^{CSR}** is a set of coordinates distributed according to the Complete Spatial Randomness (CSR) model in \mathcal{V} [6,11]:

$$F(r) = \int_0^r p_F(r) dr \tag{2}$$

The calculation of functions *G* and *F* are not computationally demanding even though function *F* requires the generation of set X^{CSR} . *G* is better suited to characterize short scale properties of particle clusters, while *F* characterizes the space devoid of particles and large scale organization of particle clusters. In order to combine their advantages, function *J* was defined as:

$$J(r) = \frac{1 - G(r)}{1 - F(r)}$$
(3)

However, the practical applications of function *J* are limited because it is not defined for F(r) = 1 and numerical problems appear when $F(r) \rightarrow 1$.

2.1.2. Monovariate second-order analysis

Ripley's functions *K*, *L* and *O* are used for second-order analysis. Because cellular compartments (VOIs) typically have an irregular shape and many particles are found close to the compartment borders, it is not possible to determine edge-corrections for Ripley's functions analytically [6,15]. Therefore, we proceeded to implement Ripley's function where edge-corrections are performed numerically, for each particle separately. Specifically, function *L* is obtained by the linearization of Ripley's *K* function [6], which facilitates the interpretation. For a set of particles located at positions \mathbf{x}_i within a 3D VOI of arbitrary shape \mathcal{V} , Ripley's functions *K* and *L* are defined as follows:

$$K(r) = \frac{4\pi r^3}{3} \cdot \frac{\sum_{i=0}^{n} C(\mathbf{x}, \mathcal{S}_i(\mathbf{x}, r))}{\lambda \cdot \sum_{i=0}^{n} V(\mathcal{S}_i(\mathbf{x}, r))} \qquad L(r) = \sqrt[3]{\frac{3K(r)}{4\pi} - r}$$
(4)

 $C(\mathbf{x}, S_L(\mathbf{x}_i, r))$ is the number of particles located in the neighborhood $S_L(\mathbf{x}_i, r)$ and $V(S_L(\mathbf{x}_i, r))$ is the volume of this neighborhood, $C(\mathbf{x}, S)$ and V(S) are estimated numerically as described in Sections 2.2.2 and 2.2.3 respectively. Neighborhood $S_L(\mathbf{x}_i, r)$ is defined as the edge-corrected neighborhood of the particle located at \mathbf{x}_i , which is obtained by the intersection of the VOI and the spherical neighborhood of the particle at radius r:

$$\mathcal{S}_{L}(\mathbf{x}_{i}, r) = \{ \forall \mathbf{x} \in \mathcal{V} | d(\mathbf{x}, \mathbf{x}_{i}) \le r \}$$
(5)

where *d* is a distance metric (Fig. 1A). In an unbounded space $\mathcal{V} = \mathbb{R}^3$, or for particles located far from boundaries, $\mathcal{S}_L(\mathbf{x}_i, r)$ is simply a sphere centered at \mathbf{x}_i with radius *r*.

Ripley's function O is defined as:

$$O(r) = \frac{\sum_{i=0}^{n} C(\mathbf{x}, S_0(\mathbf{x}_i, r, \Delta r))}{\sum_{i=0}^{n} V(S_0(\mathbf{x}_i, r, \Delta r))}$$
(6)

Here, $S_0(\mathbf{x}_i, r, \Delta r)$ is defined as the edge-corrected shell-like neighborhood around the particle located at \mathbf{x}_i , which is obtained by the intersection of the VOI and the spherical shell or radius r and thickness Δr centered at the particle location (Fig. 1B):

$$\mathcal{S}_0(\mathbf{x}_i, r, \Delta r) = \{ \forall \mathbf{x} \in \mathcal{V} | r - \Delta r/2 \le d(\mathbf{x}, \mathbf{x}_i) \le r + \Delta r/2 \}$$
(7)

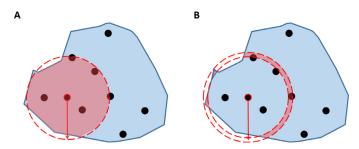


Fig. 1. Particle neighborhood and shell for Ripley's functions. (A) Edge corrected particle neighborhood S_L used for Ripley's *L* is formed as the intersection of the particle spherical neighborhood (shown as the red circle) and VOI (blue), and is indicated by light red areas. (B) Edge corrected particle shell S_0 used for Ripley's *O* is formed as the intersection of the particle spherical shell (shown as the red ring) and VOI (blue), and is indicated by light red areas. In both cases particles are represented by black points. Shown in 2D for clarity. (For interpretation of the references to colour in this figure legend, the reader is referred to the web version of this article.)

Similarly, in Eq. (6), $C(\mathbf{x}, S_0(\mathbf{x}_i, r, \Delta r))$ is the number of particles located in the shell $S_0(\mathbf{x}_i, r, \Delta r)$ and $V(S_0(\mathbf{x}_i, r, \Delta r))$ is the volume of this shell. Therefore, our definitions of $S_0(\mathbf{x}_i, r)$ and $S_0(\mathbf{x}_i, r, \Delta r)$ provide edge corrections for functions *L* and *O*.

Despite their conceptual similarity, these functions show important differences. Function O(r) can be considered more precise because it depends only on particle pairs having distance close to r, while K(r) and L(r) receive contributions from all scales between 0 and r. However, because O(r) can be understood as a derivative of K(r) it is more affected by noise, especially when the density of particles is low, which limits its usefulness in practice.

2.1.3. Bivariate analysis

The bivariate PPA functions allow investigations of the relationships between two different particle patterns, the reference pattern, $\mathbf{X}^r = \{\mathbf{x}_i^r \in \mathcal{V} | i = (1, ..., n)\}$, and the evaluation pattern, $\mathbf{X}^e = \{\mathbf{x}_i^e \in \mathcal{V} | i = (1, ..., n_e)\}$.

The distribution function *G* for the bivariate analysis is based on the nearest distances among particles between the two different patterns, $p_G^{re}(r)$:

$$G^{re}(r) = \int_0^r p_G^{re}(r) dr$$
(8)

where $p_{fe}^{re}(r)$ are distances from all reference particles to their nearest evaluation particles. There is no bivariate counterpart for function *F* and consequently neither for *J*.

The bivariate versions of the Ripley's second-order functions analysis show the co-localization of the evaluation particles in respect to the reference particles. These functions are very similar to their monovariate counterparts Eqs. (4) and (6):

$$K^{re}(r) = \frac{4\pi r^3}{3} \cdot \frac{\sum_{i=0}^{n} C^e(\mathbf{x}^e, \mathcal{S}_{L}(\mathbf{x}^r_{i}, r))}{\lambda^e \cdot \sum_{i=0}^{n} V(\mathcal{S}_{L}(\mathbf{x}^r_{i}, r))} \qquad L^{re}(r) = \sqrt[3]{\frac{K^{re}(r)}{4\pi} - r}$$
(9)

$$O^{re}(r) = \frac{\sum_{i=0}^{n} C^{e}(\mathbf{x}^{e}, \mathcal{S}_{0}(\mathbf{x}^{r}_{i}, r, \Delta r))}{\sum_{i=0}^{n} V(\mathcal{S}_{0}(\mathbf{x}^{r}_{i}, r, \Delta r))}$$
(10)

except that C^e is the number of evaluation particles \mathbf{X}^e within the edge corrected neighborhood or the shell of reference particles \mathbf{X}^r .

2.1.4. Null-models and statistical inference

In order to determine the statistical significance of particle distribution analysis obtained by the first and second order PPA functions, statistical hypothesis testing methods are used to compare the experimental with the null-model results. Consequently, when generating proper null-models for cryo-ET data, the shape of the particles and cellular compartments has to be considered. The CSR model was previously defined for point particles located in arbitrary regions [6]. Here we introduce the Complete Spatial Randomness with Volume exclusion model (CSRV) as an extension of CSR that takes into consideration the 3D shape of particles by imposing volume exclusion to avoid particle overlap. Statistical comparison of experimental results with CSRV null-model allows discarding the random particle distribution hypothesis (the null hypothesis), in which case it can be concluded that the particle organization is controlled by a structural process. Furthermore, this comparison can show whether the experimental distribution is more clustered or more uniformly distributed than CSRV and determine length scale(s) at which the differences are found. A more detailed analysis may require a null-model specifically designed for the actual experimental question [2].

To assess the statistical significance of the results obtained by the first order PPA functions *G* and *F*, the Kolmogorov-Smirnov (K-S) test is used. It requires that the variable studied is continuous, and it is applied to the PPA functions represented as the cumulative frequency distribution functions [11] to determine $D_{n,m}$ in the following way:

$$D_{n,m} = \sup_{r_D} |\tilde{G}_n(r) - G_m(r)| \tag{11}$$

 $G_n(r)$ is *G* or *F* function, obtained from experimental data comprising *n* particles and $G_m(r)$ the CSRV null-model function, obtained by simulating *m* synthetic instances of the null-model with the same number of particles and the same VOI as the experimental dataset. The null hypothesis stating that experimental and the simulated distributions are identical can be rejected with the probability $1 - \alpha$ if [20]:

$$D_{n,m} > \sqrt{-\left(\frac{m+1}{2nm}\right)\ln\left(\frac{\alpha}{2}\right)}$$
(12)

The K-S test determines a single significance value of a given first order PPA function, by taking into account the entire range of distances that forms its domain. When testing function *G*, the positive sign of $\tilde{G}_n(r_D) - G_m(r_D)$ specifies that the experimental particles are clustered and the negative sign signifies that they are uniformly distributed. The interpretation is the opposite when testing function *F*.

For the second order PPA functions, the null hypothesis can be tested by ranking PPA results obtained for multiple null-model simulations and constructing an interval of confidence (IC) as a function of distance *r*:

$$IC(r) = \left[IC(r)^{-}, \ IC(r)^{+} \right] = \left[L_{\alpha}(r), \ L_{1-\alpha}(r) \right]$$
(13)

where $L_{\alpha}(r)$ is the value of the second order PPA function under consideration that has the rank of $\alpha \cdot 100\%$. If an experimental PPA function falls outside of the interval *IC*(*r*) at the distance *r*, we can reject the null hypothesis with a confidence $1 - \alpha$. Additionally, if the experimental estimator is greater (smaller) than *IC*(*r*)⁺, we can conclude that the pattern is more clustered (more uniformly distributed) than the null-model.

2.2. Methods

2.2.1. Computation of the second-order PPA functions

Efficient implementations of many linear algebra operations that are required to compute the PPA functions introduced above are already available in libraries written in different programming languages. Here we present implementations of the operations that are currently not available in the standard libraries.

The central part of the second-order function computations concerns the determination of the number of particles $C(\cdot)$ and the volume $V(\cdot)$ of local edge corrected neighborhoods S_L and S_0 .

Because the neighborhoods are defined in a VOI of arbitrary geometry \mathcal{V} , these computations cannot rely on closed formulas but require numerical approximations. Therefore, they require computations of kernels for the number of particles and volume for neighborhoods of different radii, for each particle \mathbf{x}_i separately: $C(\mathbf{x}, \mathcal{S}_L(\mathbf{x}_i, r)), C(\mathbf{x}, \mathcal{S}_O(\mathbf{x}_i, r, \triangle r)), V(\mathcal{S}_L(\mathbf{x}_i, r))$ and $V(\mathcal{S}_O(\mathbf{x}_i, r, \triangle r))$. As a consequence, second-order analysis is computationally much more demanding than first-order analysis. We note that the global particle density can be computed by $\lambda = n/V(\mathcal{V})$.

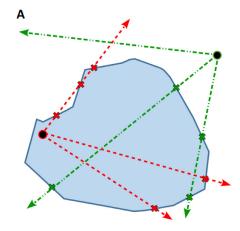
2.2.2. Counting the number of particles in a region of arbitrary shape

It is critical to choose the most efficient way of evaluating whether a point belongs to a VOI ($\mathbf{x} \in \mathcal{V}$) because this condition is evaluated intensively during the computation of second-order functions. It is required both for the computations involving real data and the randomly generated data used for null-models. Here we consider two approaches that use different ways to represent VOI.

In the surface approach, VOI is represented as a closed surface defined by a triangle mesh. To evaluate the condition $\mathbf{x}_i \in \mathcal{V}$, where \mathbf{x}_i is the particle center, we used the ray-firing method consisting of the following steps: (1) Randomly oriented rays originating at the evaluation point are generated. (2) For each ray, the number of intersections with the bounding surface is counted. (3) If the number of rays having an odd number of intersections is larger than the number of rays having an even umber of intersections, the point is considered to belong to the VOI (Fig. 2A).

In the second approach, a VOI is represented by *true* voxels in a dense 3D Boolean array (termed 3D array representation). A point belongs to a VOI simply if it belongs to the set of *true* voxels (Fig. 2B).

It is clear that the surface approach uses less memory than the 3D array approach, even though the array is binary. However, checking the condition $\mathbf{x} \in \mathcal{V}$ for the 3D arrays approach has the complexity O(1), while in the surface approach the complexity is $O(N_{rays})$, where N_{rays} is the number of rays used by the iterative ray-firing method [21]. Our results showed that for both approaches the running time per particle decreased with the number of particles (Fig. S1). This is likely a consequence of the hierarchical cache architecture of current processors. Importantly, for up to 1000 particles, the running time was smaller for the 3D array approach. Only for more than 1000 particles did the running time of the surface approach became comparable to that of the 3D array approach. In addition, the running times were similar for the cases when particles were randomly distributed in a sphere or on a spherical shell.



2.2.3. Estimating the volume of particle neighborhoods

To calculate the volume of an edge corrected neighborhood V(S), we propose two different algorithms. These parallel the approaches to determine whether a point belongs to a region introduced in the previous section.

Monte Carlo surface algorithm (MCS) requires the surface VOI representation. The volume of an edge-corrected neighborhood is calculated by randomly generating points within a spherical neighborhood, using the surface approach to determine whether these points belongs to VOI and counting the number of points inside the VOI (Alg. S1, Fig. 3A).

The second, the 3D array algorithm, requires the 3D array VOI representation. The volume of an edge-corrected neighborhood is determined by counting the number of voxels that belong to both the spherical neighborhood and the VOI using the array approach (Alg. S4, Fig. 3B). The 3D array used here has to be large enough to hold the entire neighborhood S for the largest neighbor radius r. To speed up the computations, the distance between each particle and all other voxels is pre-computed, so that a single distance calculation for a particle can be used for all neighborhoods of that particle.

To evaluate the precision of the neighborhood volume estimation by MCS and 3D array algorithms, we applied them to particles located at the distance of 5 nm to the boundary of a rectangular VOI, and also calculated the edge corrected neighborhood volumes analytically. In this way, particle distance to the VOI was small compared to the size of the VOI (200 nm) and the spherical neighborhood maximum radius (80 nm). We use a rectangular VOI to be able to determine the ground truth analytically. The volume estimation error *E* is defined as:

$$E[\%] = \frac{\hat{V} - V}{V} \cdot 100 \tag{14}$$

where *V* is the ground truth volume (computed analytically) and \hat{V} the estimated volume.

Our data showed that for both spherical and shell edge corrected neighborhoods, 10 - 1000 points and 10 rays per point the MCS algorithm performed well for all neighborhood radii (Fig. S2). As expected, increasing the number of points decreased the variability.

The 3D array algorithm was also very precise, with noticeable errors only for very small neighborhoods, (Fig. S3). This algorithm is expected to be less precise for small and non-flat neighborhoods because in such conditions the discretization of space by voxels deviates from the most the real (curved) shapes.

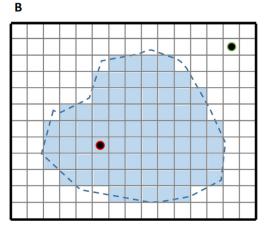


Fig. 2. Counting particles in a region of arbitrary shape. A) The surface approach: VOI (blue field) is stored as a triangle mesh (blue line), and the number of intersections ("x" symbols") between rays (dashed lines) and the VOI surface are counted. B) The 3D array approach: VOI is represented as *true* voxels (blue field). VOI boundary is shown as dashed blue line. (For interpretation of the references to colour in this figure legend, the reader is referred to the web version of this article.)

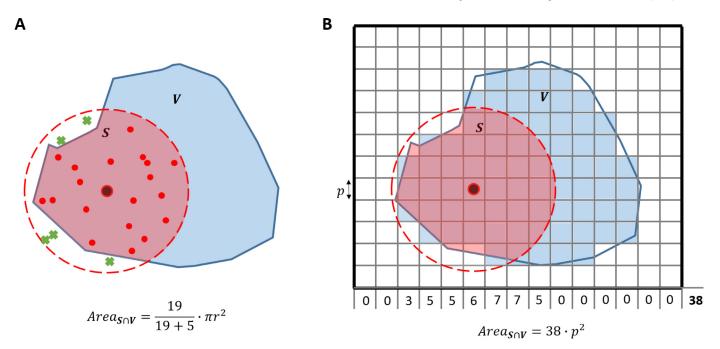


Fig. 3. 2D schematics for irregularly bounded area computation. A) MCS algorithm. B) 3D array method. Both panels show VOI (blue area), a particle (small black circle), the spherical neighborhood (red dashed circle) and the edge corrected neighborhood of the particle (red area). Ray-originating points in A) that contribute to the volume are shown as small red circles and those that do not contribute as green x-symbols. (For interpretation of the references to colour in this figure legend, the reader is referred to the web version of this article.)

In comparison, the MCS algorithm required around 1000 particles to reach the precision of the 3D array algorithm. Another advantage of the 3D array algorithm is its non-stochastic nature.

MCS required a larger running time for neighbor size up to 75 voxels and was comparable to 3D array for larger sizes (Fig S4). Here we also used 10 rays per point for MCS. In both algorithms, the running time per volume or area unit decreased with the increased neighborhood size.

2.2.4. Particle overlap

Particles encountered in cryo-ET are not points but have a finite, possibly complex 3D shape. Therefore when a new particle is added to a synthetic null-model instance, it is necessary to ensure that the particle does not overlap with any other particle. Here we use the VTK library [21] to implement volume exclusion between particles. Namely, we first use this library to generate the particle surface at the intended position and orientation specified by Euler angles. Next, we check whether the bounding box of the particle overlaps with any of the other particle bounding boxes. If it does, we use again the iterative ray-firing method implemented in VTK to ensure that no point of a particle surface is inside the surface of any previously inserted particle. If the particles indeed overlap, the new particle is rejected. This procedure ensures that particles do not overlap in 3D.

2.2.5. Parallel implementation

Second-order metric implementations are very CPU intensive. They require long running times for real data comprising dozens of tomograms and thousands of particles each. Moreover, to achieve a sufficient statistical confidence, more than 100 simulations per tomogram should be computed. To solve both problems, we provide a parallel implementation based on the multiprocessing package of the Python programming language, which uses a shared memory environment and exploits the internal parallelism of current multi-core processors.

Specifically, particles are processed separately by n independent *computational units*. To compute the second-order functions for real data, these units are evenly distributed over $p \le n$ processes that

share access to VOI. For each particle, the distance map (Euclidean or geodesic, see Section 3.1.2), is computed once and used for all distances, which minimizes the number of times a distance map has to be computed (Fig. S5A, B). For simulated data, the workload is divided by the number of tomograms to be simulated (*m*) which are executed by $p \le m$ processes (Fig. S5C).

The computational speedup [22] obtained by parallelization was almost linear up to 15 processes and continued to increase until the maximum number of processes (35) for both real and simulated data (Fig. S5D). For a high number of processes, the speedup was a little higher for the simulated data. This is beneficial for our applications because most of the workload is generated by the analysis of the null-model. Five synthetic tomograms with 500x500x100 voxels and 200 particles each were used for second-order function computations. To ensure a fair comparison, the number of simulated tomograms on each instance was the same as the number of concurrent processes.

2.2.6. Computational requirements

This software package has been developed in Python and is available open-source (see Code availability), in order to facilitate its dissemination in the research community and the development of future extensions. Graphs are plotted using *matplotlib* library [23], and surface meshes are stored and processed using VTK [21] and visualized through Paraview [24]. All computing experiments were executed on a computer node of the Max Planck Institute of Biochemistry cluster, it has 500GB RAM with 36 processors Intel(R) Xeon(R) CPU E5-2699 v3 @ 2.30GHz with SUSE Linux Enterprise Server 12 SLES 12 SP 1 Operation System.

3. Results

3.1. Validation with synthetic data

3.1.1. Uni- and bivariate functions

To validate the implementation of the second-order metrics, we generate synthetic tomograms that contain randomly distributed and clustered particles (Fig. 4A, B). To simulate clustered particles,

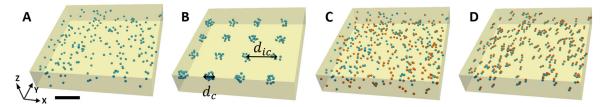


Fig. 4. Synthetic datasets used for validations. A) CSRV. B) SRPV with cluster diameter $d_c \approx 46$ voxels and intercluster distance $d_{ic} = 125$. C) Two independently generated CSRV patterns (blue and red). D) CSRV pattern (blue), points of the spatially correlated pattern (red) are located at distances obtained from the normal distribution $\mathcal{N}(\mu = 10, \sigma = 1)$ to the CSRV pattern, we chose the parameters of the normal distribution so that the two patterns visually colocalize. In all cases particles (blue and red) were spheres of radius $r_p = 5$ each particle pattern contained 200 particles and tomograms had size 500x500x100 voxels. Scale bar 100 voxels. (For interpretation of the references to colour in this figure legend, the reader is referred to the web version of this article.)

we introduce here the Sinusoidal Random Pattern (SRP) distribution \tilde{U} :

$$\tilde{U}((x = q\pi x', y = q\pi y', z = q\pi z') | \sin x + \sin y + \sin z > 3t)$$
(15)

where x', y' and z' are random real numbers uniformly distributed in the interval [-1, 1], $q \in N$ and $t \in [0, 1]$. Distribution \tilde{U} creates clusters that have oval shape of maximal radius $r_c \approx (q\pi)^{-1} \arccos(3t-2)$ (in x', y', z' units). The clusters reach an almost spherical shape for t > 0.8, in which case the radius can be approximated by $r_c \approx (q\pi)^{-1} \sqrt{6(1-t)}$. The distance between neighboring clusters (center to center) is $d_{ic} = 2/q$ (in x', y', z' units).

Here we generated multiple CSRV, and SRP with volume exclusion (SRPV) synthetic datasets, where SRPV was created from SRP populated with finite size particles (as opposed to point-particles) and excluding overlapping particles. Our software allows using any closed surface to represent particles. For simplicity, here we used spheres with radius $r_p = 5$. This synthetic datasets had a size of 500x500x100 voxels, where 500 voxels corresponds to x', y', and z' interval [-1, 1]. For SRPV, we set q = 4, and t = 0.8, thus generating 16 clusters at the inter-cluster distance of $d_{ic} = 125$ voxels (Fig. 4B).

To validate our implementation of the univariate Ripley's functions, we first checked the *L* function obtained for the CSRV dataset (the null-model). The mean value of 100 simulations was close to 0 and the *IC* 5 – 95% was spread around the 0 value, as expected for a pure random pattern (Fig. 5A, D). This was the case for the whole distance range except for distances lower than $2r_p$, indicating that at short distances particle volume exclusion dominates. We also calculated the *L* function for additional five CSRV simulations, their mean shows the stochasticity expected for the analysis of real randomly distributed particles.

The univariate Ripley's functions that we obtained for 100 SRPV simulations did not differ between MCS and the 3D array algorithms, Next, we calculated the univariate Ripley's functions L and O functions for the clustered, SRPV, datasets (Fig. 5B, C, E, F). Using the MCS and the 3D array algorithms produced virtually indistinguishable results, as was the case for the CSRV null-model (compare panels A-C and D-F of Fig. 5). The calculated L and O functions for SRPV model showed significant deviations from the results obtained for the CSRV null-model. Specifically, the distance at which function L reached maximum and function O minimum were close to $d_c/2$ and d_c respectively, as expected based on analytical calculations [25]. Furthermore, function L was zero and its first derivative was positive at the distance that corresponds to the inter-cluster distance. This can be explained by observing that the concentration of particles in a neighborhood of radius that equals the periodicity of the point pattern is simply the global particle concentration. At the same distance, function O reached a local maximum, which directly points to clustering at that distance d_{ic} .

Finally, to validate bivariate Ripley's functions, we generated datasets containing two particle sets, the particle sets were spa-

tially uncorrelated in some and spatially correlated with each other in other datasets. The uncorrelated datasets consisted of two independent CSRV patterns (Fig 4C). To generate the correlated datasets, we first generated a CSRV pattern and then placed particles of the second set at distances following the Normal distribution $\mathcal{N}(\mu, \sigma)$ from randomly selected CSRV particles (Figs. 4D and 6.A). For both MCS and the 3D array algorithms, the mean value of the bivariate L function for 100 simulations of the nullmodel was very close to 0 at all distances, except for the very short ones where the effects of volume exclusion dominate, thus showing that the two sets were independent (Fig. 6B). For the correlated datasets, the bivariate L function of the correlated dataset was significantly different from those obtained for the uncorrelated nullmodel at an intermediate range of distances. The bivariate L function for the correlated dataset reached significance at the distance of approximately $\mu - \sigma$ and reached a maximum at $\mu + \sigma$, which agrees with the criterion customarily used in the field [25].

All together, our implementation of the second-order uni- and bivariate functions on synthetic datasets yielded results that allowed the determination of the correct clustering distance scales. Furthermore, the MCS and the 3D array based implementations produced almost identical results, thus validating both algorithms.

3.1.2. Influence of the distance metric

Because a VOI can have an arbitrary 3D shape and it may include holes, the choice of distance metric *d* used for defining the neighborhoods S_L and S_0 , Euclidean or geodesic, can have a strong impact on the computation of the second-order functions. We used the Distance Transform (DT) [26] for computing Euclidean and Fast Marching Method (FMM) algorithm [27] for calculating geodesic distance. We implemented FMM only for the 3D array VOI representation.

For a VOI that has convex shape and trivial topology in 3D (such as that shown on Fig. 6A), Euclidean and geodesic distances are the same, resulting in essentially the same L and O functions (Fig. 6B, C).

However, cellular VOIs often have a complex shape because they are delineated by biological membranes or are formed by a particular distribution of molecular complexes. To evaluate the PPA functions obtained using the two distance metrics, we generated a synthetic dataset where VOI takes the shape of a cropped spherical shell of radius $r_m = 50$ and thickness $t_m = 6$ voxels (Fig. 7A). Such VOIs are encountered for complexes bound to a membrane of a cellular organelle [28]. The particle pattern was formed by four clusters localized in the VOI, and for each cluster *i*, particles were distributed according to the following expression:

$$U_{i}(x = x_{c} + r_{m}\cos(\phi + \varphi_{i})\sin\phi,$$

$$y = y_{c} + r_{m}\sin(\phi + \varphi_{i})\sin\phi, z = z_{c} + r_{m}\cos\theta)$$
(16)

where $c_m = (x_c, y_c, z_c)$ are the coordinates of the VOI, $\mathcal{N}(\mu = 0, \sigma = 0.25)$, r_m is the radius of the VOI, θ and ϕ are spherical angles taken from the normal distribution and $\varphi_{1,2,3,4} =$

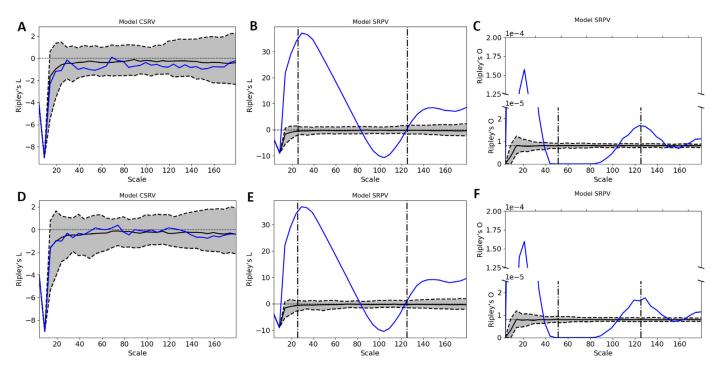


Fig. 5. Univariate second-order metrics validation with synthetic datasets. (A-C) 3D array and (D-F) MCS algorithm. (A, D) Function *L* for CSRV model, blue line shows the mean of additional five CSRV simulations. (B, E) Function *L* for SRPV model, shown are the mean of five simulations (blue line) and the CSRV null-model (grey). (C, F) Function *O* for SPRV model, shown are the mean of five simulations (blue line) and the CSRV model, shown are the mean of 100 CSRV simulations, respectively. For MCS the number of iterations for convergence is 1000 and the maximum 100000. In (B, E) the vertical dashed lines mark $d_c/2$ and d_{i_c} , and in (C, F) d_c and d_{i_c} respectively. In (A, B, D, E) the horizontal line marks X = 0, and in (C, F) the global density λ , in all cases these lines represent the behavior of the ideal random pattern. (For interpretation of the references to colour in this figure legend, the reader is referred to the web version of this article.)

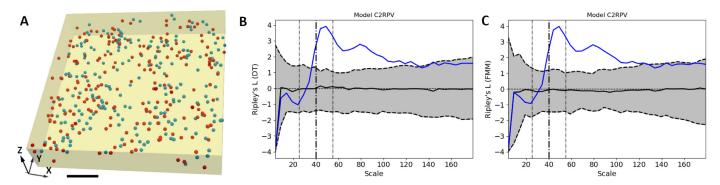


Fig. 6. Bivariate second-order analysis validation. (A) Two correlated point patterns where each particle of pattern 2 (red) are placed at a distance controlled by a distribution $\mathcal{N}(\mu = 40, \sigma = 5)$ to a particle of pattern 1 (blue). (B) Function L for the correlated pattern shown in A (mean of 5 simulations) is shown in blue, and the *IC* 5 – 95% of 100 simulated uncorrelated pairs in gray. Distances were measured using Distance Transform (DT). (C) Like B except that Fast Marching Method (FMM) was used for distance computation. (B, C) Thick vertical lines represent $r = \mu$ and the thin ones $r = \mu \pm 3\sigma$. Tomograms size 500x500x100. Scale bar 100 voxels. (For interpretation of the references to colour in this figure legend, the reader is referred to the web version of this article.)

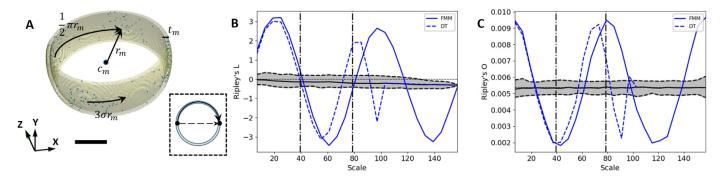


Fig. 7. Geodesic and Euclidean distance based second order functions on a complex VOI. (A) A biological membrane-like VOI, $r_m = 50$, and $\sigma = 0.25$ containing clustered particles (blue points). The inset shows the difference between geodesic (solid arrow) and Euclidean distances (dashed arrow) between two points on the opposite sides of a sphere. (B) Function *L* and (C) function *O* for the clustered particle pattern (blue lines) and the null-model (*IC* 5 – 95% in gray, black line represents the mean). In both cases the second order functions based on geodesic distance (solid and blue line) and Euclidean distance (dashed blue line) are shown. Scale bar 30 voxels. (For interpretation of the references to colour in this figure legend, the reader is referred to the web version of this article.)

 $\{0, \pi/2, \pi, 3\pi/2\}$ define cluster centers. In this way, the centerto-center inter-cluster distance was $\frac{1}{2}\pi r_m$ (78.5 voxels for r_m of 50 nm) and the cluster size can be approximated by $3\sigma r_m$ (37.5 voxels). In addition, particle distribution $\tilde{U}(x, y, z)$ takes volume exclusion into account.

Our results showed that functions L and O were significantly different from the CSRV null-model. The geodesic distance based L function had the first maximum at approximately 20 voxels $(1.6\sigma r_m)$, a 0 crossing with a negative slope at 40 voxels and another with a positive slope at 80 voxels, while the geodesic O function had a local maximum at 80 voxels (Fig. 7B, C). These correspond to the size of the clusters and the distance between them, in the same way that functions L and O did for a simple VOI (Figs 4A and 5B, C, E, F). However, the Euclidean distance based L and O functions showed a shift towards shorter distances. Additionally, the geodesic distance based O function of the null-model particle pattern (CSRV), was almost perfectly flat, as expected for the ideal case (Fig. 7C). The L function for the null-model deviated slightly from the expected 0-value at the largest distances, corresponding to the neighborhoods that reach to the opposite side of the VOI (Fig. 7B). This is likely because the geodesic distance between some of the distant points on the cropped spherical shell VOI differs from the geodesic distance on a full spherical shell. Therefore, our results show that using FMM allows an accurate determination of the geodesic distance and that for membrane-like compartments and other complex shape VOIs, the PPA functions should be calculated based on geodesic and not Euclidean distances for.

3.2. Experimental data

We proceeded to validate the computational methods described in previous sections on experimental data. To this end, we applied the first and second order PPA functions to study the organization of ribosomes in yeast cells visualized by cryo-ET. We used data form two experimental groups, each comprising a set of tomograms from yeast cell cytoplasm imaged in situ. The first set contains 13 tomograms of yeast cells that were treated with Rapamycin and the second 14 tomograms of untreated cells [29]. In both cases, cell cytoplasm was segmented, ribosomes were localized by template matching and a high resolution structure of ribosomes (obtained from EMD-3068) was used to represent the ribosome shape, as shown on a slice of a labeled tomogram (Fig. 8A). These masks are created manually, or first automatically and then corrected manually, which ensures they are close to if not completely accurate. Considering that the mask volume is $> 10^6$ nm³ and that only individual voxels may be mislabeled (segmentation voxel size 2.096 nm), the accuracy level can be estimated to be >99%. In the example shown in (Fig. 8A), particle localization was determined by template matching [29]. The precision error was below 2.096 nm because template matching was done at every voxel. While the segmentation and localization accuracies in general depend on the tomogram quality and the segmentation and localization methods used, the accuracy levels stated above are reached routinely.

We calculated the first-order functions for each tomogram separately in order to determine the lengths that characterize the organization of ribosomes (Fig. 8.B-E). In untreated cells, the nearest neighbor function showed a clear peak that represents the most common nearest-ribosome distance. For each tomogram, multiple null-model distributions were generated, each having the same VOI and the same number of ribosomes as the corresponding tomogram. All functions (*G*, *F*, and *J*) showed that ribosomes are significantly clustered with respect to the null-model (CSRV). This is in agreement with the expected aggregation of functional ribosomes to form poly-ribosomes. More generally, this type of analysis can help describing macromolecular interactions [2,17]. However, it is not straightforward to combine the first order functions of individual tomograms within an experimental group to obtain a single function that characterizes an experimental condition, because the particle concentration within VOIs differ between the cells even though they were grown under identical conditions, which influences the first order functions (Fig. 8F-I).

Second-order PPA functions were calculated for each tomogram separately, as well as for their respective CSRV null-models (Fig. 9). The null-model *L* functions for different tomograms and different experimental groups (control and Rapamycin treated) were very similar. The experimental group means were almost indistinguishable, only the variability was higher for the Rapamycin set likely because of the lower number of particles (Fig. 9.A). In contrast, function *O* for the null-model was very different for different tomograms and there was a clear separation between the two experimental groups. Because this variability was likely due to the different global particle concentration in the tomograms, we here propose to use the radial distribution function g(r) [30], which is computed by normalizing function *O* by the global particle concentration:

$$g(r) = \frac{1}{\lambda}O(r) \tag{17}$$

This normalization restored the low variability of the null-model simulations between tomograms and experimental groups (see Fig. 9.C).

When applied to the experimental data, the *L* and the radial distribution functions clearly showed significant clustering, both for all tomograms taken together and for the experimental groups taken separately (Fig. 9.D, F). In both experimental cases the first maximum was located around 25 nm, which approximately corresponds to the double of the most frequent nearest neighbor distance Fig. 8.B). While the large variability between the groups obtained for the function *O* precludes the interpretation of all tomograms taken together, each of the experimental group showed significant clustering when compared to the corresponding null-model (Fig. 9B, E). While it is expected that in the untreated cells the formation of polyribosomes leads to ribosome clustering, our results show that some form of ribosome clustering persists in the Rapamycin treated cells.

Furthermore, in order to set the stage for a statistical comparison between the experimental groups, we computed the mean and IC 5 - 95% as we did before, except that because of the low number of tomograms, the *IC* contains all tomograms. This comparison is valid for *L* and radial distribution function as null-models converge to a similar *IC*. However, in our case, this approach can not be applied to function *O*, because we already saw that it is sensitive to particle global densities and the two experimental groups contain different number of particles. The statistical significance between the groups can be easily established for distances at which the *L* and the radial distribution functions show clear separation between the experimental groups. To determine the significance at other distances, non-parameteric inference tests could be applied to values of the second order functions.

Considering that there is no particle overlap, we estimate that possible errors arising from imperfect segmentation and localization are well below the errors caused by the volume estimation (Figs S2 and S3). In order to confirm this, we recalculated 2nd order functions with randomly shifted protein position using the Normal distribution $\mathcal{N}(\mu = 0, \sigma = 1)$. The results obtained (Fig. S6) are virtually indistinguishable from those presented in Fig. 9.D,F.

4. Discussion

Here we implemented the first and second-order, mono- and bivariate PPA functions for 3D VOIs of arbitrary shape. This is of

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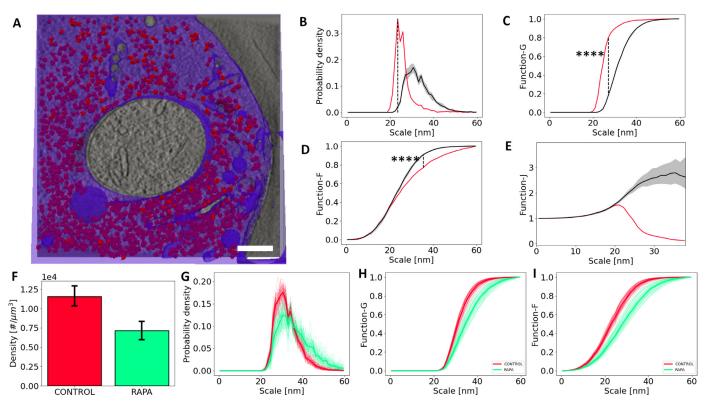


Fig. 8. First-order analysis of experimental data. A) Tomographic slice of the analyzed cryo-tomograms, with overlay showing segmented cytoplasm (VOI, transparent blue) and localized ribosomes (red), scale bar 200 nm. B) Histogram of all nearest neighbor distances (function *G*), dashed vertical line marks the most frequent nearest neighbor distance. C) Function *G* shown as a cumulative distribution (K-S test, $D_{2182,20} = 0.6182$ and $\alpha < 0.0001$). D) Spatial contact distribution function (*F*) with 1000 simulated points (K-S test, $D_{1000,20} = -0.1389$ and $\alpha < 0.0001$). E) Function *J*. B-E) The functions obtained for tomogram showed in A), experimental data are shown in red. CSRV null-model simulation means are shown in black and *I*C 5 – 95% in grey (20 simulations). Dashed vertical lines show $D_{n.m}$. F-I) Comparison between the two sets of tomograms, control (CONTROL, red), and Rapamycin treated (RAPA, green). F) Particle density within the VOIs, mean and *I*C 5 – 95%. G) Function *G* shown as thistogram. H) Function *G* shown as the cumulative distribution. I) Function *F*. G-I) Analyses of individual simulated null-models are shown as thin lines, thick lines show the means. In all cases Euclidean distance was used. (For interpretation of the references to colour in this figure legend, the reader is referred to the web version of this article.)

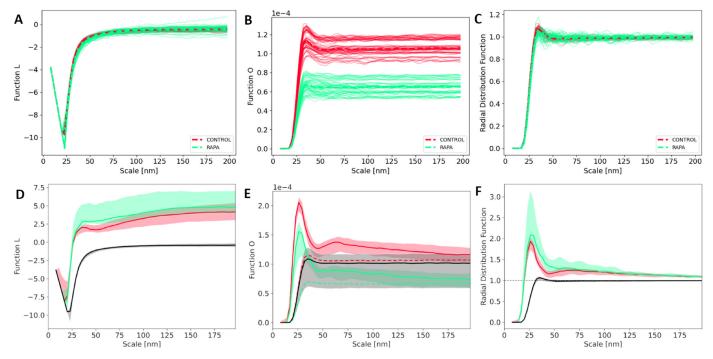


Fig. 9. Second-order analysis for experimental and the corresponding CSRV null-model data. A-C) Functions obtained for CSRV null-model data corresponding to the control (CONTROL, red) and Rapamycin treated (RAPA, green) tomograms. Each thin semi-transparent line represents a CSRV null-model simulation, five per tomogram, the thick dashed lines show the means. D-F) Functions obtained for experimental tomograms, control (CONTROL, red) and Rapamycin treated (RAPA, green). Pooled null-model simulations (untreated and Rapamycin treated) IC 5 - 95% is shown in grey and the mean in black. E) Also shows the mean simulations of the experimental groups (dashed lines). A, D) Function *L*. B, E) Function *O*. C, F) Radial distribution function. (For interpretation of the references to colour in this figure legend, the reader is referred to the web version of this article.)

particular importance for the analysis of molecular complexes visualized in 3D biological images, such as those obtained by cryo-ET of cellular samples, because cellular regions often have complex, non-convex shape. These cellular regions can be be formed by lipid membrane-bounded organelles and vesicules, membranes where certain membrane-bound complexes reside, or any other cellular regions that constrain the localization of molecular complexes of interest. Furthermore, because molecular complexes occupy a substantial part of the cellular volume and the influence of their size cannot be neglected at short distances, we represented molecular complexes as 3D objects and imposed volume exclusion to prevent the overlap.

The implementation of the second-order functions was particularly important because they are sensitive to clustering at multiple distance scales, thus providing more information than the first order functions. It critically depends on solving two tasks; (1) determination of the number of particles located within an irregularly bounded spatial region and (2) measuring the volume of this space. The solutions depend on the approach taken to represent spatial regions (particle neighborhoods and VOI). In the first approach, we represented spatial regions by a triangular surface mesh that defines the region boundary. Here we used the stochastic ray-firing algorithm to count the number of particles and MCS for the volume determination. The second approach we implemented is based on representing spatial regions as 3D binary arrays, which makes the implementation of the two tasks trivial. In theory, ray-firing and MCS are stochastic methods that can achieve any precision and are independent of the volume shape. 3D array based methods are deterministic and they typically use more memory than the surface based methods. However, our results show that for the same computational time, the 3D arraybased methods achieved a higher precision than the surface-based counterparts. Nevertheless, the second order PPA functions that we computed using these two approaches were virtually indistinguishable.

We validated the implementation of the second-order, uni- and bi-variate PPA functions on synthetic datasets. Our software correctly detected the clustering distance scales that characterized the particle distribution in the synthetic datasets. Additionally, we implemented the PPA functions based on both Euclidean and geodesic distances, verified that they produced the correct results, and showed that for simple VOIs the Euclidean and the geodesic results were the same. Importantly, the geodesic distance based second order PPA functions were more suitable for applications involving complex VOIs, such as for describing the nanodomain organization of molecular complexes located on curved membranes, the situation commonly encountered in cellular environments.

Because analytical solutions for the PPA functions do not exist for complex VOI and particle geometries, synthetic random tomograms (null-models) are required in order to determine the statistical significance of the PPA functions applied to experimental data. The inherent variability of cellular components and the fact that a typical cellular cryo-ET dataset contains tens of tomograms necessitate generating a set of synthetic random model tomograms for each experimental tomograms. An experimental tomogram and its corresponding synthetic tomograms have to have the same VOI and the number of particles. Furthermore, a large number of synthetic tomograms is needed to allow reaching a specified significance level. To alleviate the computational burden involved in the generation of the necessary number of random models and the computationally intensive calculation of the PPA functions for all tomograms, we propose a multi-process implementation of these routines that reduced the running times, thus enabling the effective analysis of realistic datasets. We achieved a speed-up factor of approximately 15 using a single processor multi-core architecture, a similar value to those recently obtained in [31] where a cluster of computers was used for computing space-time Ripley's *K*.

Application of our software to yeast cell cytoplasm imaged by cryo-ET allowed us to detect ribosome clustering and determine the characteristic distance scales. Furthermore, we showed that the second order functions, in particular L and the radial distribution functions, were better suited to compare experimental groups comprising multiple tomograms because they were not sensitive to the global concentration of particles.

Future applications of our software are also expected to provide a spatial characterization of macromolecular crowding, as well as liquid and lipid phase separation. These processes recently gained a significant biological interest, because they were shown to affect biochemical reactions in cells and global organization of cellular membranes and regions in different cellular systems [32,33].

Therefore, our implementation of the PPA functions provides a tool that can characterize simultaneous clustering at multiple distance scales, which is suitable for applications to cellular molecular complexes visualized by cryo-ET, as well as to other 3D systems where real-size particles are located within regions possessing complex geometry.

Code availability

PyOrg code belongs to a more general public repository (pyseg_system, https://github.com/anmartinezs/pyseg_system) which contains additional software, tests, and dependencies. The code of PyOrg alone can be found within code/pyorg folder. https://github. com/anmartinezs/pyseg_system/tree/master/pyseg-sys/pyorg.

Declaration of Competing Interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

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Supplementary material

Supplementary material associated with this article can be found, in the online version, at doi:10.1016/j.cmpb.2022.106693.

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