

Seasonality of North Atlantic phytoplankton from space: impact of environmental forcing on a changing phenology (1998–2012)

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Abstract

Seasonal pulses of phytoplankton drive seasonal cycles of carbon fixation and particle sedimentation, and might condition recruitment success in many exploited species. Taking advantage of long-term series of remotely sensed chlorophyll *a* (1998–2012), we analysed changes in phytoplankton seasonality in the North Atlantic Ocean. Phytoplankton phenology was analysed based on a probabilistic characterization of bloom incidence. This approach allowed us to detect changes in the prevalence of different seasonal cycles and, at the same time, to estimate bloom timing and magnitude taking into account uncertainty in bloom detection. Deviations between different sensors stressed the importance of a prolonged overlap between successive missions to ensure a correct assessment of phenological changes, as well as the advantage of semi-analytical chlorophyll algorithms over empirical ones to reduce biases. Earlier and more intense blooms were detected in the subpolar Atlantic, while advanced blooms of less magnitude were common in the Subtropical gyre. In the temperate North Atlantic, spring blooms advanced their timing and decreased in magnitude, whereas fall blooms delayed and increased their intensity. At the same time, the prevalence of locations with a single autumn/winter bloom or with a bimodal seasonal cycle increased, in consonance with a poleward expansion of subtropical conditions. Changes in bloom timing and magnitude presented a clear signature of environmental factors, especially wind forcing, although changes on incident photosynthetically active radiation and sea surface temperature were also important depending on latitude. Trends in bloom magnitude matched changes in mean chlorophyll *a* during the study period, suggesting that seasonal peaks drive long-term trends in chlorophyll *a* concentration. Our results link changes in North Atlantic climate with recent trends in the phenology of phytoplankton, suggesting an intensification of these impacts in the near future.

Keywords: MODIS, North Atlantic Ocean, Ocean Colour, Phytoplankton phenology, Remote sensing, SeaWiFS

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Introduction

Seasonal pulses of phytoplankton growth set the rhythm of marine ecosystems (Barnes & Hughes, 1999), and represent an important influx of atmospheric CO₂ into the oceans (Falkowski *et al.*, 1998). These events drive seasonal cycles of particle sedimentation (Deuser & Ross, 1980; Honjo, 1982; Billett *et al.*, 1983) and are tightly linked to the ecology of zooplankton (Longhurst, 2007), including the early stages of many exploited species (Hjort, 1914; Cushing, 1990; Townsend *et al.*, 1994; Durant *et al.*, 2007). The timing and characteristics of seasonal peaks are a major indicator of the functioning of marine pelagic ecosystems (Platt & Sathyendranath, 2008; Racault *et al.*, 2012). In land, both the phenology of vegetation and migratory species have been altered by recent climate change (Peñuelas & Filella, 2001;

Parmesan, 2007; Sletzer & Post, 2009), while the sea analyses based on long-term field sampling programmes have shown consistent changes both in the phenology and biomass of marine plankton (Reid *et al.*, 1998; Edwards *et al.*, 2001; Edwards & Richardson, 2004).

The ephemeral nature of changes in phytoplankton and their spatial extent make their characterization difficult by classical sampling techniques. This has been remediated to some extent by the availability of decade long, high-quality remotely sensed monitoring of chlorophyll *a* concentration (hereafter, chl *a*) (McClain *et al.*, 2004a; McClain, 2009). Analyses incorporating satellite data have revealed a tight link between climate variability and recent decreases in phytoplankton biomass and primary productivity at the global scale (Gregg & Conkright, 2002; Antoine *et al.*, 2005; Gregg *et al.*, 2005; Behrenfeld *et al.*, 2006; Martinez *et al.*, 2009; Vantrepotte & Mélin, 2009), the expansion of low chl *a* concentration areas in the subtropics (McClain *et al.*, 2004b; Polovina *et al.*, 2008; Irwin & Oliver, 2009) and a

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1 decline in mean phytoplankton cell size (Polovina &
2 Woodworth, 2012). Studies on marine phenology have
3 focused on the main peak of phytoplankton growth in
4 temperate and polar regions, i.e. the spring phytoplank-
5 ton bloom, and have highlighted the great variability in
6 this event and a trend towards an early occurrence of
7 these blooms in northern latitudes in recent years (Siegel
8 *et al.*, 2002; Platt & Sathyendranath, 2008; Henson *et al.*,
9 2009; Kahru *et al.*, 2011; Zhai *et al.*, 2011), as well as the
10 importance of trophic mismatches (Beaugrand *et al.*,
11 2003; Platt *et al.*, 2003; Koeller *et al.*, 2009; Kristiansen
12 *et al.*, 2011). Secondary pulses during the fall in temper-
13 ate latitudes and autumn/winter blooms in subtropical
14 and tropical regions have received in general less atten-
15 tion (but see Ueyama & Monger, 2005; Martinez *et al.*,
16 2011; Cole *et al.*, 2012; Sapiano *et al.*, 2012).

17 Here, we combine SeaWiFS and MODIS data to study
18 recent changes in the seasonality of phytoplankton in
19 the North Atlantic Ocean. We develop a methodology
20 which accommodates the different nature of spring and
21 autumn/winter blooms, allowing us to study both
22 events simultaneously. At the same time, the method
23 allows propagating uncertainty in bloom detection to
24 estimates of the change in the extent of areas presenting
25 different types of seasonal cycles, as well as to estimates
26 of the timing and magnitude of seasonal peaks. On the
27 basis of this approach, we examine whether recent
28 ocean colour observations reveal (i) a geographical shift
29 in the incidence of different kind of seasonal cycles; (ii)
30 changes in the timing and magnitude of spring and
31 autumn/winter blooms, and their relationship to
32 trends in chl *a* concentration; and (iii) the potential abil-
33 ity of different environmental factors to explain recent
34 changes in the phytoplankton seasonality in the North
35 Atlantic.

37 Materials and methods

39 Data sources and data preparation

41 A box between 110°W 10°S and 50°E 80°N was selected to
42 study changes in the seasonality of phytoplankton in the
43 North Atlantic Ocean and its marginal seas. Chlorophyll *a*
44 is commonly used as an index of phytoplankton biomass and
45 thus of changes in phytoplankton abundance or size. The
46 main advantage of chl *a* is that its concentration in the near
47 surface can be readily measured from space (McClain, 2009),
48 but at the cost of ignoring deep chlorophyll maxima. The use
49 of chl *a* as an index of phytoplankton biomass is further con-
50 founded in general by changes in nutrient availability and in
51 the light regime that modulate pigment cell levels (Laws &
52 Bannister, 1980), problems that might be especially important
53 in subtropical latitudes (see below). Moreover, changes in
54 phytoplankton species composition might alter as well the
relationship between chl *a* concentration and biomass.

Daily time series of remotely sensed chl *a* concentration
[mg m⁻³] between September 1997 and April 2013 were
retrieved from Level 3 (geolocated, corrected and averaged
over a regular grid) SeaWiFS (Sept. 1997–Dec. 2007, reprocess-
ing R2010.0) and Aqua MODIS (Jul. 2002–Apr. 2013, reprocess-
ing R2013.0) standard mapped images (SMI) available at
the Ocean Color Web (Feldman & McClain, 2012; Goddard
Space Flight Center, NASA; oceancolor.gsfc.nasa.gov). Note
that long data gaps due to instrument failures beginning in
January 2008 prevented the use of SeaWiFS data available up
to December 2010. Original data at a nominal scale of 9 km
were averaged over a 0.25°grid (cell side ca. 25 km). We used
chl *a* concentration maps estimated using the Garver-Siegel-
Maritorena semi-analytical model (GSM, Garver & Siegel,
1997; Maritorena *et al.*, 2002). The GSM presents some advan-
tages over other algorithms when data from different missions
are combined, given that it is based on a common parameteriza-
tion independent of the sensor employed to measure ocean
colour (Maritorena *et al.*, 2010; note that problems reported in
this article related to the drift of the 412 and 443 nm bands of
Aqua MODIS were corrected in the last reprocessing [R2013.0];
see Meister *et al.*, 2012 and oceancolor.gsfc.nasa.gov/WIKI/
OCReproc2013%282e%290MA.html). Despite this advantage,
deviations between different sensors are still expected as a
consequence of differences in their radiometry (Maritorena
et al., 2010). To assess the impact of our choice of the GSM algo-
rithm, we repeated all the analyses using chl *a* estimates
retrieved using the sixth version (OCv6, oceancolor.gsfc.nasa.gov/REPROCESSING/R2009/ocv6) of the OC4
(SeaWiFS) and OC3M (Aqua MODIS) empirical band-ratio
algorithms (O'Reilly *et al.*, 2000), to check the robustness of our
approach to the algorithm employed to estimate chl *a* concentra-
tion (see the Supporting Text in the Supporting Informa-
tion).

Sea surface temperature (SST) data were used to delimit
different seasons and thus to help in the detection and charac-
terization of increases in chl *a* concentration (a 5 day filter was
previously applied to time series to avoid spikes). Data for
other environmental variables and indexes were retrieved to
study the physical forcing on the timing and magnitude of
seasonal peaks during the study period. Climate variables
gridded to the same spatial and temporal scale of the chl *a*
observations included SST, incident photosynthetically avail-
able radiation (PAR), wind speed and eddy kinetic energy
(Table 1).

Characterization of seasonal changes in chlorophyll *a* concentration

Seasonal cycles of chl *a* concentration present a wide variation
in the North Atlantic, reflecting changes in physical, chemical
and biological conditions from the equator to the poles (e.g.,
Longhurst, 2007). This includes seasonal regimes character-
ized by either one or two peaks in chl *a*, which have been typi-
cally associated with seasonal changes in stratification (i.e. the
spring phytoplankton bloom) or mixing (autumn/winter
blooms) (Dutkiewicz *et al.*, 2001). Seasonal increases in chl *a*
concentration near the surface reflect both changes in

Table 1 Environmental variables employed to assess the potential importance of climate forcing to explain interannual changes in phytoplankton seasonality

Variable name (abbreviation [units])	Source and processing details	Comments
Sea Surface Temperature (SST [K])	NOAA Optimum interpolation 0.25° daily SST analysis (OISST version 2, Reynolds <i>et al.</i> , 2007). Database produced and maintained by C. Liu and R. W. Reynolds at NCDC, www.ncdc.noaa.gov/oa/climate/research/sst/oi-daily-information.php	SST not only accelerates the rate of physiological and ecological processes (of phytoplankton and of grazers, Townsend <i>et al.</i> , 1994) but it is also a tracer of vertical mixing and of the advection of waters with different properties. Warmer (cooler) waters might be related to increased (decreased) stratification and light exposure and reduced (increased) nutrient availability.
Integrated Photosynthetically Available Radiation (PAR [Einstein $\text{m}^{-2} \text{day}^{-1}$])	Daily time series of Level 3 PAR from 400 to 700 nm, available at the Ocean Color Web (Feldman & McClain, 2012; Goddard Space Flight Center, NASA; oceancolor.gsfc.nasa.gov).	Together with vertical attenuation and mixing depth, incident PAR determines the subsurface light field (e.g., Platt <i>et al.</i> , 1991). In this way, a lower PAR might alter chl <i>a</i> concentration by limiting phytoplankton growth rates or by increasing pigment cell levels, and <i>vice versa</i> .
Wind stress (τ [N m^{-2}])	Derived from daily wind speed [m s^{-1}] maps were integrated from the six-hourly, Level 3 Cross-Calibrated Multi-Platform Ocean Surface Wind Velocity product (CCMP, Atlas <i>et al.</i> , 2011; available at PO.DAAC, podaac.jpl.nasa.gov). The drag coefficient was estimated based on Yelland & Taylor, 1996; and Yelland <i>et al.</i> , 1998;	Wind stress is a proxy of wind surface mixing and turbulence, so increased wind stress is related to an increased mixed layer ventilation and nutrient renewal, as well as deeper phytoplankton entrainment, and, in principle, higher dilution and lower encounter rates with grazers (Irigoin <i>et al.</i> , 2005; Behrenfeld, 2010). Nevertheless, air–sea heat fluxes and vertical convection are also important drivers of mixed layer depth at high latitudes.
Eddy kinetic energy (EKE [$\text{m}^2 \text{s}^{-2}$])	Derived from the reference series of daily geostrophic velocity anomalies produced by Ssalto/Duacs and distributed by Aviso (www.aviso.oceanobs.com), with support from CNES (Centre National d'Etudes Spatiales). The reference series is obtained by merging data from various missions (Topex/Poseidon, Jason-1, European Remote Sensing satellites [ERS 1 and 2], and Envisat) using the methods developed by Le Traon <i>et al.</i> (1998).	Eddy kinetic energy is a proxy of variability in ocean currents and mesoscale features which might promote an early stratification and enhance bloom development (Karrasch <i>et al.</i> , 1996).

phytoplankton abundance and in the amount of chl *a* per cell. Cell concentration might change due to population growth, horizontal advection and dispersion, or as a consequence of changes in vertical distribution (Behrenfeld, 2010). Pigment cell levels vary depending on nutrient availability and on the light field (Laws & Bannister, 1980). The photoacclimation response is especially important following autumn mixing in subtropical latitudes, when the recirculation of phytoplankters in a deeper mixed layer decreases light exposure and results in an increase in chl *a* concentration in the water column (Leterrier *et al.*, 1993; DuRand *et al.*, 2001; Behrenfeld *et al.*, 2005; Siegel *et al.*, 2005; Westberry *et al.*, 2008; Vantrepotte *et al.*, 2011).

Episodic changes in surface chl *a* concentration due to other processes alter these seasonal cycles, reflecting for instance the influence of mesoscale features, coastal upwelling and land inputs, to name a few of them. Satellite data allow the identification of all these events, with the main constraint arising from limited data availability during cloudy conditions. Here we adopted a rather practical approach for the identification

and characterization of seasonal peaks, trying to avoid the influence of high frequency events. Seasonal extremes in SST were used to delimit each season and to identify candidate periods for seasonal peaks of increase in chl *a* concentration. We considered a period centred on each calendar year but covering the time period between previous and next year SST maxima (see Jönsson & Eklundh, 2002). The chl *a* time series for this period was then smoothed by fitting a generalized linear model assuming Gamma distributed errors using the canonical, inverse link function:

$$p(\text{chl}_t | \widehat{\text{chl}}_t) \sim \text{Gamma}(u, v) \\ (\widehat{\text{chl}}_t)^{-1} \sim \text{Normal}(\beta X, \sigma_\epsilon) \quad (1)$$

The estimation of the shape and rate parameters of the Gamma distribution (u and v) is surpassed in this way by linking the expected chl *a* values to the linear predictor ($\text{eta}_t = \beta X$). The linear predictor included an intercept, a linear trend on time and sine and cosine waves to represent the seasonal cycle

1 by annual and semi-annual harmonics ($\omega = 1/365$; see Vargas
2 **1** *et al.*, 2008; Sapiano *et al.*, 2012), yielding the equation:

$$3 \eta_t = \beta_0 + \beta_1 t + \beta_2 \sin(2\pi\omega t) + \beta_3 \cos(2\pi\omega t) + \beta_4 \sin(4\pi\omega t) \\ 4 + \beta_5 \cos(4\pi\omega t) + \beta_6 t \sin(2\pi\omega t) + \beta_7 t \cos(2\pi\omega t) \quad (2)$$

5 Note that the model allowed also a linear trend in the
6 amplitude of the annual harmonics. This model specification
7 was redundant for some of the series, so we determined an
8 optimal structure based on model ranks determined using the
9 Akaike Information Criterion (AIC, Burnham & Anderson,
10 2003), a relative measurement of goodness of fit that includes
11 a penalty term to weight down model complexity and avoid
12 over fitting. The model was fitted under a Bayesian frame-
13 work, employing the modified Expectation-Maximization
14 algorithm developed by Gelman *et al.* (2008) and implemented
15 in the R package *arm* (Gelman *et al.*, 2009; see also Gelman &
16 Hill, 2007). We assumed standard, weakly informative priors
17 for each j parameter in Eqn (2), i.e.:

$$18 \beta_j \sim \text{Cauchy}(\mu, v) \\ 19 \mu = 0 \text{ and } \begin{cases} v = 10, & j = 0 \\ v = \frac{2.5}{2\text{sd}(x_j)}, & j > 0 \end{cases} \quad (3)$$

22 where the location parameter μ centres our prior belief about
23 the mean of posterior parameter values in zero and the scale
24 parameter was tuned depending on the standard deviation of
25 each covariate x_j (a larger variation in x_j puts more a priori
26 weight in small values of β_j).

27 Posterior parameter distributions were then used to gener-
28 ate an envelope of model realizations (1000) that was
29 employed to propagate model uncertainty to a set of bloom
30 metrics (timing and magnitude) used to characterize the sea-
31 sonal cycle. Local extremes in chl a delimited periods of accu-
32 mulation that were considered as candidate blooms if they
33 reached a level above the 60th percentile of a Gamma distribu-
34 tion fitted to chl a observations between consecutive SST
35 extremes (i.e. a minimum and a maximum or vice versa). The
36 choice of this threshold was arbitrary but helped us to reject
37 small amplitude waves. Candidate blooms were then classifi-
38 ed either as spring or autumn/winter blooms based on the
39 relative timing of bloom metrics with respect to SST extremes.
40 We considered that a candidate bloom corresponded to a
41 spring bloom if its timing and peak occurred after the seasonal
42 SST minimum but before the SST maximum. Similarly, the

timing of autumn/winter blooms must occur between the sea-
sonal SST maximum and the next minimum (avoiding thus
possible confusions with next year's spring blooms), even if
the timing of the peak occurred after the SST minimum. For a
given model realization, only the first candidate bloom meet-
ing the criteria above was retained, although in some cases all
the candidates were rejected.

Determining the occurrence of a bloom in this way does not
presuppose the development of a bloom every year and com-
pensates to some extent the lack of data during cloudy periods
(Gregg & Casey, 2007; Cole *et al.*, 2012). It is important to note
that we defined the timing of the bloom as the day when the
net rate of increase in chl a concentration attained a maximum.
The definition is similar to other approaches based on a prede-
fined threshold level, although our intention was not to deter-
mine the date of bloom initiation. Our definition also differs
from the timing of bloom onset, defined by Sverdrup (1953) as
the date when the net rate of phytoplankton increase becomes
positive. With our definition, we tried to prevent potential
measurement errors in the net rate of increase associated with
the small changes in chl a concentration at the onset of the
bloom, and due to the lack of data during cloudy periods in
northern latitudes. On the other hand, bloom magnitude was
defined as the peak chl a concentration attained during the
bloom. We also estimated mean chl a concentration during the
entire bloom (i.e. between consecutive chl a minima), but only
to complement analyses of changes in bloom timing and mag-
nitude. A set of examples have been included in Fig. 1 and a
diagram summarizing bloom determination is included in the
Supplementary Information (Fig. S6).

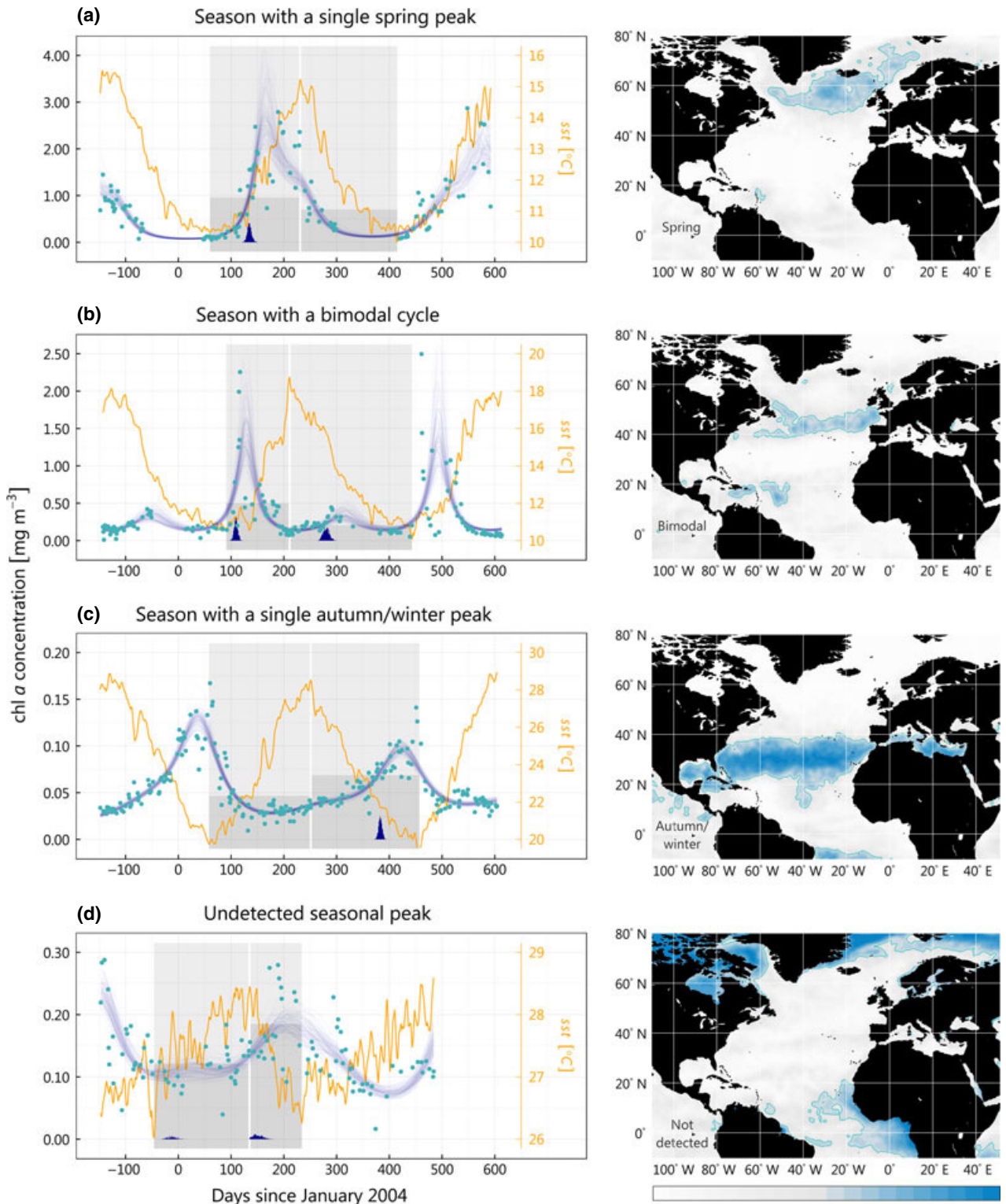
Analyses of changes in the seasonality of chl a concentration

The characterization of seasonal chl a time series resulted in
four different possibilities attending to the presence or absence
of different peaks. We distinguished mean seasonal cycles pre-
sented (i) a single spring; or (ii) a single autumn/winter
bloom; (iii) a spring bloom followed by an autumn/winter
one (i.e. a bimodal seasonal cycle); and (iv) mean seasonal
cycles in which neither of the blooms were detected. Each of
the 1000 posterior model realizations was assigned to any of
these categories, resulting in a raw estimate of the probability

43 **Fig. 1** Different types of seasonal cycles of surface chlorophyll a concentration in the North Atlantic. Rows a–d correspond to the four
44 different types of seasonal cycles distinguished. The plots on the left side of each row illustrate the approach employed to characterize
45 seasonal peaks (see also Fig. S6), while the maps on the right present the corresponding probability of occurrence of each type of sea-
46 sonal cycle in the entire North Atlantic. In the left panels, posterior simulations (blue lines) from a model fitted to remotely sensed
47 observations of chlorophyll a concentration (chl a , green dots) were used to assess the probability of detecting different peaks in phyto-
48 plankton biomass and their timing (histograms). A peak qualified as a bloom after surpassing a threshold chl a concentration (dark grey
49 shading). Identified peaks were classified either as spring or autumn/winter blooms based on the timing of sea surface temperature
50 extremes (SST, orange line), which determined spring or fall candidate periods (grey shaded areas). The series were extracted from the
51 following pixel locations: (a) 12.6°W55.1°N, (b) 9.6°W47.6°N, (c) 55.1°W30.1°N, and (d) 32.6°W0.6°S. Probability maps were derived
52 from 15 consecutive seasonal cycles (from 1998–1999 to 2012–2013; data for Aqua MODIS and SeaWiFS were averaged for overlapping
53 seasons), and were based on 1000 posterior simulations of model Eqn 2 fitted to data available during each season. The contour line
54 encloses areas with a probability of detection greater than 0.5. Fig. S8 in the Supporting Information presents maps distinguishing
regions with no data and regions where, despite data being available, no peak was detected (e.g., the case illustrated in d).

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of each kind of seasonal cycle occurring at each pixel location. These probabilities were then integrated over regions defined by grouping biogeochemical provinces delimited by Longhurst (2007) to obtain a weighted estimate of the total areal extent corresponding to each kind of seasonal cycle.

Longhurst's (2007) regions were further subset to avoid marginal seas and to account for differences in the detectability of different seasonal cycles (see Fig. S7). Trends in the prevalence of each type of seasonal cycle were analysed based on a Dirichlet regression model that included a second-order trend in

time (see Table S2 in the Supporting Information). This kind of model has an error structure ideally suited to analyse compositional data (proportions adding up to unity) in the presence of covariates (Campbell & Mosimann, 1987; Hijazi & Jernigan, 2009), and was fitted using an adaptive Metropolis algorithm (Roberts & Rosenthal, 2009).

Simulations of the model fitted to chl *a* time series were also employed to obtain an augmented sample of the timing (defined as the day of maximum net increase in chl *a*) and magnitude (defined as the peak chl *a* concentration during a bloom) of blooms occurring at each pixel location. Median timings and magnitudes were estimated at the pixel level, but detection probabilities were retained and employed to weight the reliability of different observations, propagating in this way uncertainty in bloom detection to estimates based on bloom statistics. Interannual changes in bloom metrics were analysed using a model including a fixed effects factor to account for biases between different sensors and a linear trend in time. We assumed normally distributed errors for the residuals of this model, which we considered a reasonable assumption for both bloom timing and bloom magnitude, although in the latter case only after log transformation. Temporal trends in bloom metrics were later compared to trends in mean chl *a* during the entire year, estimated after fitting Eqn. (2) to the complete daily chl *a* series using a log link function to ease interpretation. This model thus included a linear trend, a term to account for sensor bias and terms to account for a seasonal cycle potentially changing its amplitude between years.

Estimates of bloom timing and magnitude were also compared to a set of environmental factors to assess the potential importance of climate forcing to explain interannual changes in phytoplankton seasonality. Time series of linearly detrended anomalies of (i) sea surface temperature; (ii) incident photosynthetically active radiation; (iii) wind stress; and (iv) eddy kinetic energy, were standardized to mean zero and standard deviation one. All these variables modulate phytoplankton dynamics and might alter bloom timing and magnitude (Table 1). Detrended anomalies were preferred to represent short-term effects on bloom metrics and to prevent problems of collinearity in models including more than one covariate. Detrending did not affect the patterns of association found with original data in models with a single covariate. The models fitted to time series of bloom metrics assumed normally distributed errors (after log transformation in the case of bloom magnitude) and included again a linear trend and a term to account for sensor bias. Models included all the covariates, although only wind stress or one of its components was included to avoid problems of collinearity (best model structure based on AIC). The fraction of deviance explained by each environmental factor was determined by fitting models excluding sequentially each covariate.

A 0.5° spatial moving window was employed to augment areal coverage and to reduce spatial noise in all the estimates. The overlap between SeaWiFS and Aqua MODIS during four entire seasons (from 2003–04 to 2006–07 inclusive) allowed us to assess potential biases derived from using data coming from different sensors and satellites, as further detailed in the Supporting Information (see Supporting Text, Table S1 and

Figs. S1–S5). All calculations involving different pixel locations accounted for changes in cell area with latitude using the reference ellipsoid WGS84.

Results

Incidence of different types of seasonality

The cumulated probability of different types of seasonal cycle presented a clear latitudinal pattern with a single spring bloom in the subpolar Atlantic, a single autumn/winter bloom in subtropical latitudes, and a bimodal cycle in temperate latitudes (Fig. 1). There was a clear transition among each pair of regions, with a relatively sharp gradient in detection probabilities (decaying shade intensity outside green contours in Fig. 1). No single type of seasonal cycle dominated in transitional regions, although recurrent blooms were detected when seasons with single and double peaks were pooled together to estimate the probability of occurrence of spring and autumn/winter blooms (indeed they qualified for the estimation of trends in blooms statistics, Fig. 3). This explains for instance the failure to highlight spring blooms in the north-western Mediterranean (Bosc *et al.*, 2004), that were obscured due to the detection of bimodal cycles in some years. Our approach failed to detect a marked seasonal cycle in pixel locations north of $\sim 70^\circ\text{N}$ due to data scarcity (see Figs. S2 and S8 in the Supporting Information), and in most of the tropical and equatorial Atlantic, due to multimodal and highly irregular small amplitude seasonal oscillations (Longhurst, 2007). The western tropical Atlantic, near the Antilles, presented an exception to this general pattern. Bimodal and even spring seasonal cycles were common in this region where seasonal peaks are driven by enhanced biological N_2 fixation (Coles *et al.*, 2004; Subramaniam *et al.*, 2008).

The prevalence of different types of seasonal cycles changed between seasons (Fig. 2; see also Table S2). The Dirichlet regression results in nonlinear trends in prevalence, so model based estimates for the first and last years will be used to illustrate changes during the study period (1998–2012). In the polar and subpolar North Atlantic (Fig. 2a–d), the prevalence of seasons with a single spring bloom was coupled to changes in the frequency of pixel locations where no bloom was detected (Kendall's $\tau = -0.75$ [$-0.59, -0.90$]; model-based estimates of the median and 90% posterior density interval). The prevalence of spring peaks presented also a decrease (e.g., from an estimated fraction of 0.44 [0.43, 0.45] in 1998–1999 to 0.33 [0.30, 0.37] in 2012–2013, equivalent to a change in extent of 5.14 [3.28, 6.91]

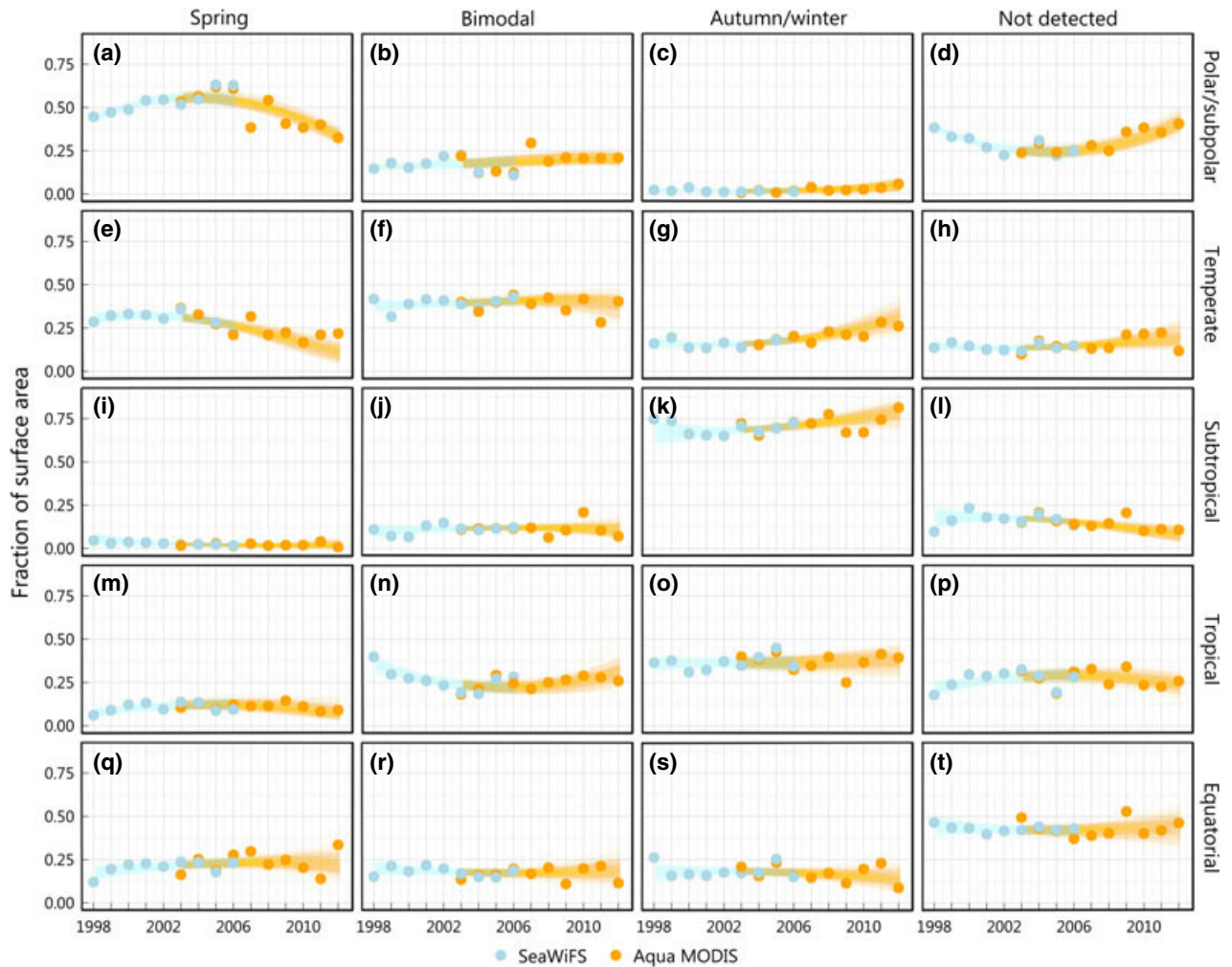


Fig. 2 Prevalence of different types of seasonal cycle in the main biogeochemical regions of the North Atlantic. Each panel presents the incidence of each type of seasonal cycle after correcting for differences in detectability between sensors. Estimates were derived from a weighted integral considering the probabilities for each kind of seasonal cycle (e.g., Fig. 1). Lines correspond to posterior simulations ($n = 400$) from a Dirichlet regression model accounting for differences between sensors and including a second degree polynomial trend to account for nonlinear time trends (Table S2). Alpha blending was employed to represent model uncertainty, with $\alpha = 1/80$ (i.e. the overlap of 80 lines correspond to full opacity). The regions group biogeochemical provinces defined by (Longhurst, 2007) (see Fig. S7 in the Supporting Information). See Fig. S9 in the Supporting Information for the same figure based on chl *a* estimates retrieved using the OCv6 band-ratio algorithm.

$\times 10^5 \text{ km}^2$), that was compensated in part by a weak increase in bimodal cycles ($2.14 [0.85, 3.85] \times 10^5 \text{ km}^2$).

In temperate regions (Fig. 2e–h), there was a clear decline in the prevalence of locations with a single spring bloom (fraction of area reduced from 0.31 [0.28, 0.34] to 0.11 [0.06, 0.16], equivalent to a reduction in extent of $13.2 [9.8, 16.0] \times 10^5 \text{ km}^2$). Although this trend was again negatively related to changes in detectability (e.g., $2.5 [-0.0, 6.1] \times 10^5 \text{ km}^2$; $\tau = -0.81 [-0.36, -1.00]$), it was compensated mainly by an increase in the extent of locations with a single autumn/winter bloom ($7.68 [3.89, 12.28] \times 10^5 \text{ km}^2$, $\tau = -0.98 [-0.79, -1.00]$) and, to a lesser extent, by an

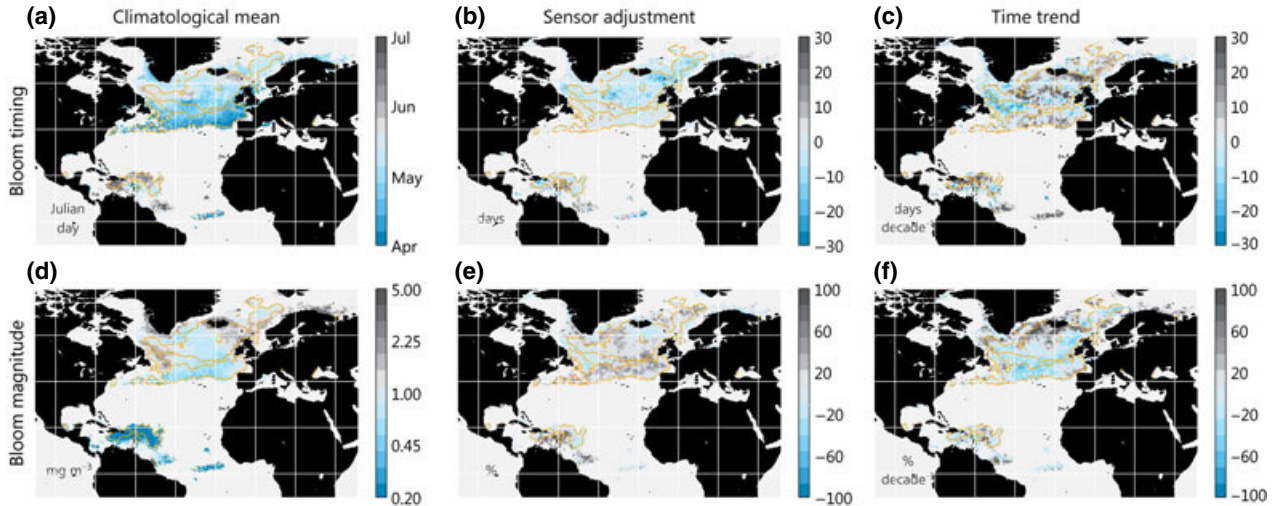
increase in bimodal cycles ($1.25 [-3.07, 4.78] \times 10^5 \text{ km}^2$, $\tau = -0.50 [0.41, -0.93]$). In the subtropics (Fig. 2i–l), posterior estimates indicated a slight increase in the dominance of seasonal cycles with a single autumn/winter cycle (from a fraction of 0.68 [0.61, 0.73] to 0.78 [0.71, 0.84]), mainly at the expense of a reduced incidence of areas where no peak was detected ($\tau = -0.93 [-0.66, -1.00]$). No clear temporal trends were detected in tropical and equatorial regions (Fig. 2m–t), where the prevalence of different seasonal cycles remained almost constant except for some high frequency excursions away from mean prevalence levels coinciding with El Niño events (e.g., 2005–06).

Timing and magnitude of seasonal peaks

Mean levels and time trends in the timing (defined as the day when the net rate of increase in chl *a* concentration was maximized during the phase of accumulation in modelled chl *a* series) and the magnitude of seasonal peaks (i.e. the peak chl *a* concentration attained)

presented a marked spatial structure during the study period (Fig. 3a,d,g,j; see also Fig. S10 and especially Fig. S13 in the Supporting Information for uncertainty associated with these estimates). The comparison of bloom statistics based on data retrieved by different sensors resulted in differences structured in space that were especially important in the case of bloom magni-

Spring bloom



Autumn/winter bloom

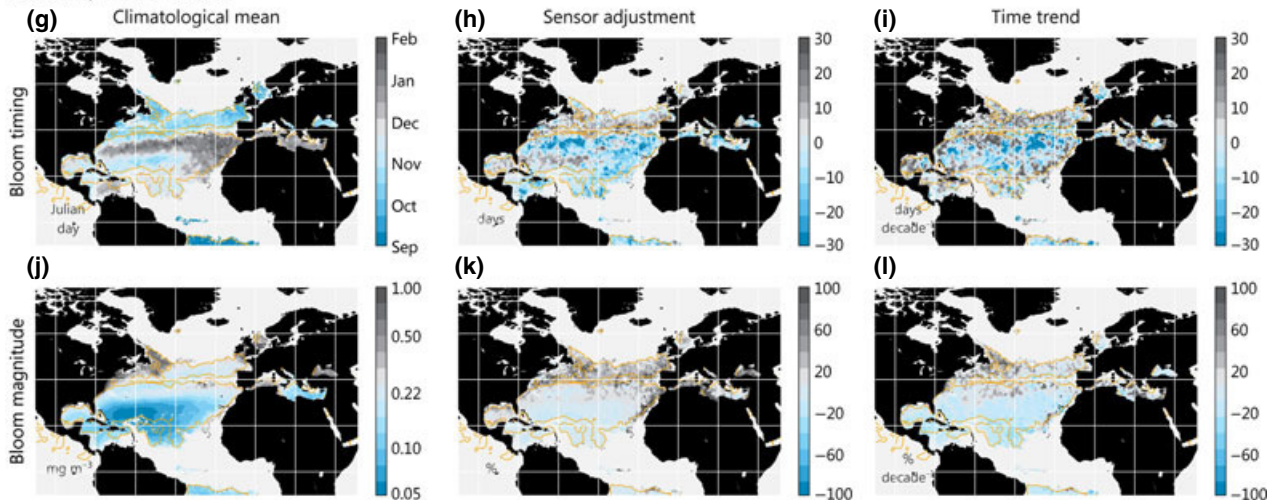


Fig. 3 Timing and magnitude of spring (*upper panels*) and autumn/winter blooms (*lower panels*). Maps in each column correspond to the posterior mean date (Julian day) and chl *a* concentration (mg m^{-3}), the differences between sensors (in days and as a percentage respectively) and the time trends (days or percentage per decade) during 1998–2012 (units are also indicated in the bottom left corner of each map). A model including a linear trend in time was fitted to estimates of bloom statistics. The model included also a fixed effect factor to account for biases in mean bloom metrics between different sensors. Bloom statistics estimated from SeaWiFS data were taken as baseline. The timing (defined as the day when the net rate of increase in chl *a* concentration attained a maximum during each wave of increase) and the magnitude (defined as the peak chl *a* concentration attained during a bloom) of spring and autumn/winter blooms was determined for each season and pixel location from posterior simulations of the smoothing model fitted to chl *a* observations (Eqn 2). Only locations where the cumulated probability of detecting either type of bloom was greater than 0.5 were considered. Estimates corresponding to single-peaked and bimodal seasonalities were pooled together. Orange contours delimit regions where the probability of each type of seasonal cycle is greater than 0.5 (see Fig. 1). Figs. S13 and S14 in the Supporting Information present the uncertainty associated with all these estimates and results based on chl *a* estimates retrieved using the OCv6 band-ratio algorithm.

tude (Fig. 3e,k). Spring peaks were detected early in April in the temperate North Atlantic and up to June in the Subpolar gyre (Fig. 3a), although there were some noticeable exceptions, especially in coastal regions. The mean magnitude of spring peaks covered almost two orders of magnitude ($0.2\text{--}14.4\text{ mg m}^{-3}$, Fig. 3d) and increased with latitude, although it was mainly influenced by the proximity to land. In the western subtropical Atlantic ($10^{\circ}\text{--}23.5^{\circ}\text{N}$), low magnitude spring peaks ($0.13\text{ [}0.06,0.65\text{] mg m}^{-3}$) occurred in late June, except in the region influenced by the Amazon river outflow (median peaks of up to 9.4 mg m^{-3} , but note that chl *a* concentration retrievals are less reliable in case 2 waters). The timing of autumn/winter peaks presented a more complicated pattern (Fig. 3g). Small amplitude seasonal peaks ($0.13\text{ [}0.06,0.31\text{] mg m}^{-3}$, Fig. 3j) were detected in late November and December in the southwestern side of the Subtropical gyre, and occurred up to early February towards the north and in the eastern side. The autumn/winter bloom of bimodal seasonal cycles of temperate regions was more intense ($0.26\text{ [}0.19,0.90\text{] mg m}^{-3}$). The timing of bimodal autumn/winter blooms occurred later towards the south; as

early as September in regions like the North Sea and up to early November near the Subtropical gyre.

The timing and the magnitude of either spring or autumn/winter blooms presented in general a positive temporal association at the pixel level, except in the case of spring blooms in some locations in the western Subpolar gyre and in polar latitudes (Fig. 4, see also Fig. S11). On a large scale, this relationship resulted in a slight predominance of regions with either delayed and more intense blooms or advancing blooms declining in magnitude (29.0 and 30.1%, respectively, i.e. 59.0% of the area presented trends of the same sign, Fig. 4). Nevertheless, trends in timing were more heterogeneous in space than trends in bloom magnitude. Trends towards delayed blooms predominated at the basin scale (58.7%; similar figures for either kind of bloom). Declining bloom magnitudes were more frequent in the case of autumn/winter blooms (62.3%), and to lesser extent, in the case of spring blooms (54.1%). The magnitude of spring blooms increased in coastal regions and in the northern North Atlantic, while the magnitude of autumn/winter blooms increased mainly in regions presenting a bimodal cycle.

COLOR

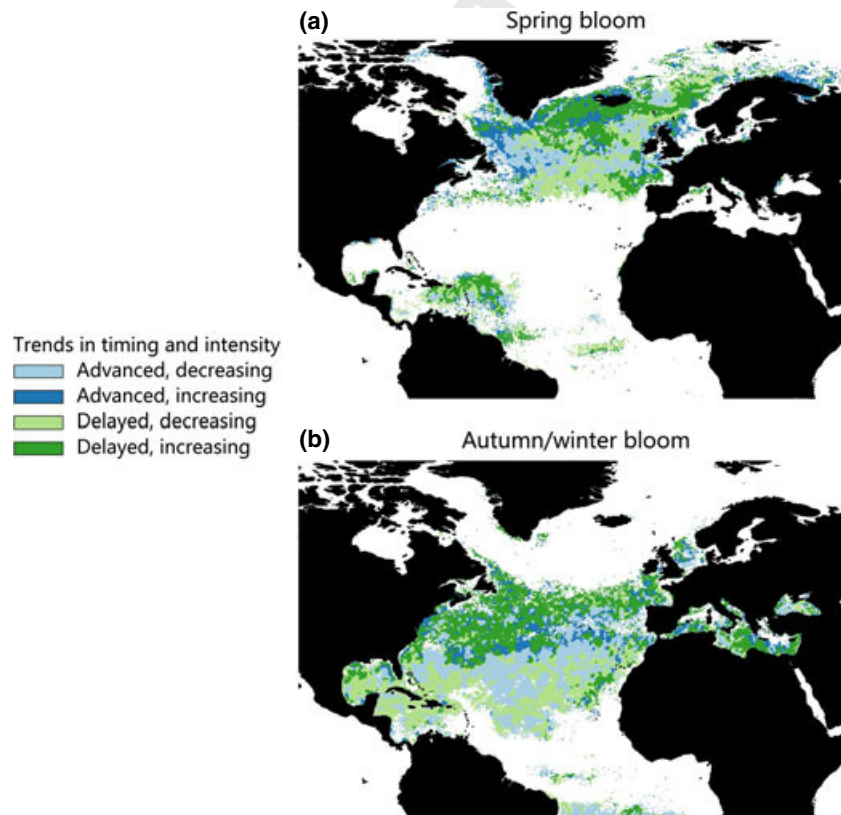


Fig. 4 Categorical maps showing the association between the sign of time trends in the timing and in the magnitude of spring and autumn/winter blooms in the North Atlantic during 1998–2012. Each colour depicts the four combinations between blooms with an advancing or delaying timing and an increasing or decreasing magnitude.

1 In general, delays and advances in the timing of
 2 spring and autumn/winter blooms were of the same
 3 magnitude, although changes in the mean magnitude
 4 of blooms of either sign were more important in the
 5 case of spring blooms (Fig. 3, see also Figs. S10 and
 6 S15). Advanced and less intense spring blooms were
 7 common in polar regions (north of 65°N) and in both
 8 sides of the Atlantic between 45°N and 60°N. In con-
 9 trast, trends towards delayed and more intense blooms
 10 were common in the Faroe-Iceland ridge, in the Irming-
 11 er Sea and, in general, in areas north of 55°N in the
 12 central North Atlantic (e.g., 65.7% of the locations
 13 within the box 40°W 55°N and 20°W 65°N). Delayed
 14 spring and autumn/winter blooms also predominated
 15 along the southern limit of regions presenting bimodal
 16 cycles in temperate latitudes. Interestingly, trends
 17 towards a longer interval between the timing of both
 18 blooms predominated in temperate regions presenting
 19 bimodal cycles (71.4%). At the same time, the magni-
 20 tude of spring blooms declined whereas autumn/winter
 21 blooms were more pronounced. This last result
 22 contrasted with the conspicuous predominance of
 23 trends towards less intense blooms in regions with a
 24 single autumn/winter bloom (67.1%). Advanced
 25 autumn/winter blooms predominated in the interior of
 26 the Subtropical gyre, although delayed blooms were
 27 common along its southern boundary and in marginal
 28 seas like the Mediterranean Sea and the Caribbean Sea.

29 Mean chl *a* concentration decreased in most of the
 30 North Atlantic during the study period (60.5%). This
 31 pattern was determined mainly by the higher preva-
 32 lence of negative trends in the tropical and subtropical
 33 North Atlantic (Fig. S17). Changes in mean chl *a* con-
 34 centration varied between regions depending on the type
 35 of mean seasonal cycle. Areas with a single autumn/
 36 winter bloom presented in general a decrease in mean
 37 chl *a* (78.0%), while increasing and decreasing trends
 38 were equally important in areas with bimodal seasonal
 39 cycles (46.5%). In the case of areas with a single spring
 40 bloom, trends towards increased chl *a* predominated
 41 (81.2%). This contrasting response is indicative of the
 42 tight link between changes in seasonal peaks and dec-
 43 adal changes in mean chl *a* concentration. We further
 44 examined the correspondence between changes in
 45 bloom timing and magnitude and changes in mean chl
 46 *a* concentration at the pixel level (Fig. S18). Changes in
 47 the magnitude of seasonal peaks were of the same signs
 48 as changes in mean chl *a* (i.e. coherent changes in 59.2
 49 and 77.4% of the areas presenting spring and autumn/
 50 winter blooms respectively), something relevant consid-
 51 ering the lower coherence with trends in mean chl *a*
 52 during the entire bloom (53.4 and 64.0%, i.e. between
 53 consecutive chl *a* minima). On the other hand, no clear
 54 association was detected between the sign of trends in

bloom timing and in mean chl *a* concentration (i.e.
 coherent sign of trends in just 49.3 and 53.4% for spring
 and autumn/winter blooms respectively).

Impact of environmental factors on the timing and magnitude of blooms

The influence of environmental factors on spring and
 autumn/winter blooms presented a similar spatial pat-
 tern for bloom timing and magnitude (Figs. 4, S19 and
 S20). Each environmental variable presented marked
 and spatially coherent trends that in some cases varied
 between the cold and the warm seasons (Fig. S21).
 These trends also varied spatially, although they were
 indicative of a basin scale trend towards warmer sur-
 face waters and increased cloudiness (i.e. reduced
 PAR). Changes in wind stress were more complex, with
 a decline in most of the basin associated with the nega-
 tive trend in the North Atlantic Oscillation index dur-
 ing the study period (Henson *et al.*, 2009; Hurrell &
 Deser, 2009), but with increased wind stress in some
 locations, especially along the western North Atlantic.
 After removing these trends, models containing all the
 environmental factors considered at the same time –
 but including only total wind stress or one of its com-
 ponents to avoid collinearity problems – explained an
 amount of variation in bloom timing and magnitude of
 0.24 [0.08,0.62] (fraction of deviance explained, median
 and 90% interval, Fig. S19). The importance of different
 variables varied depending on latitude and on the type
 of bloom (Fig. S20). The variable contributing to explain
 more deviance in bloom metrics presented a patchy dis-
 tribution, although changes in wind forcing played a
 critical role in determining bloom characteristics in
 most of the basin (Fig. 5).

Sea surface temperature (SST) during the cold season
 was positively associated with delayed and more
 intense blooms in the polar North Atlantic and in the
 Irming er Sea (Fig. S20a). This result contrasted with the
 prevailing negative association found in the rest of the
 North Atlantic, either when considering spring or
 autumn/winter blooms. Zonal wind stress (τ_u) and, to
 a lesser extent incident PAR, appeared as the most
 important factors in the rest of the northern North
 Atlantic (Fig. 5). Seasons with stronger westerlies cor-
 responded to delayed and more intense spring blooms,
 while seasons with PAR above the average presented
 delayed and, to less extent, enhanced spring blooms
 (Fig. S20e and b respectively). The positive association
 between a delay in the timing of the bloom and its mag-
 nitude might seem counterintuitive, although it is par-
 tially a consequence of the way we defined the timing
 of the bloom. If we assume that chl *a* dynamics obey a
 logistic function, an increase in the carrying capacity of

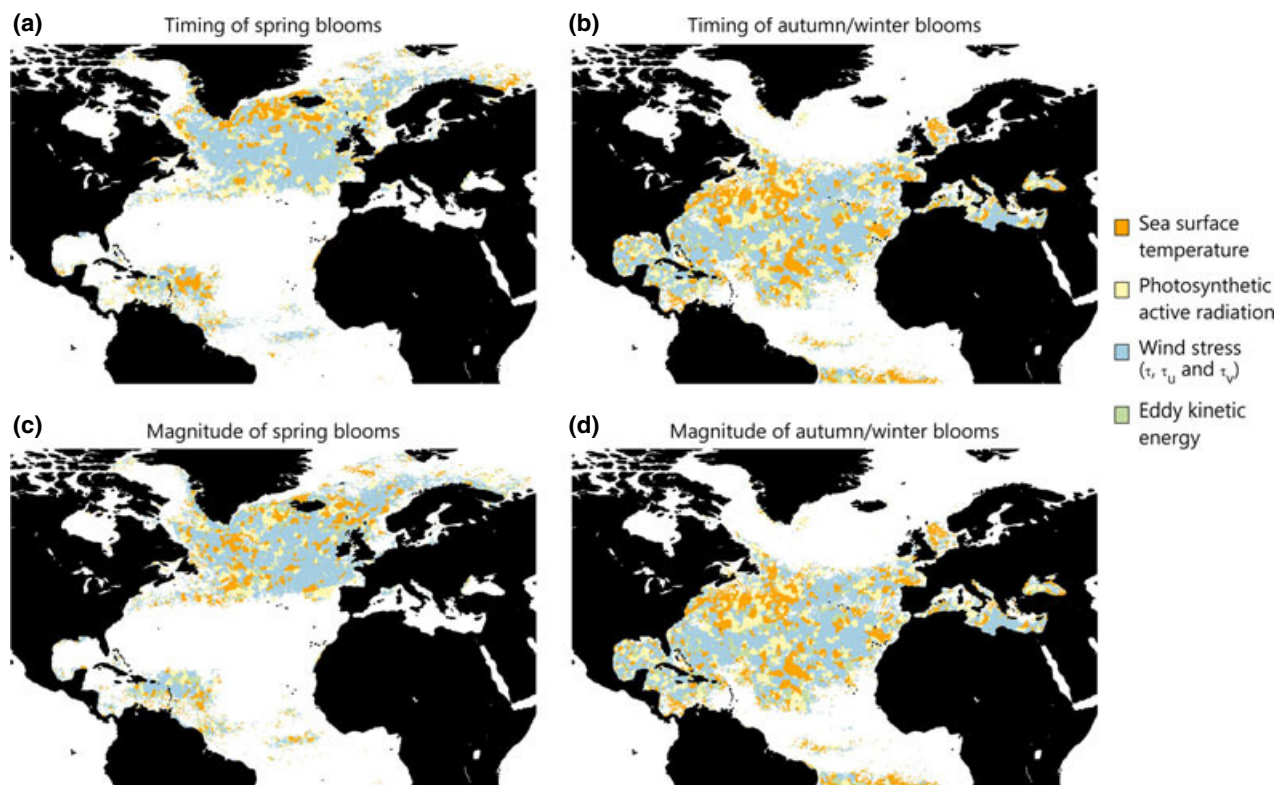


Fig. 5 Environmental variables explaining most deviance in interannual changes in the timing (*upper maps*) and magnitude (*lower maps*) of spring and autumn/winter blooms in the North Atlantic during 1998–2012. A model including all the covariates considered in the study (Table 1) was fitted to bloom statistics, although combinations resulting in problems of collinearity were excluded (e.g., those including wind stress and its components). The covariate explaining more deviance was determined by comparing the decline in deviance explained after deleting each covariate one at a time. The amount of deviance explained is presented in the Supplementary Information.

the environment might cause a delay in the peak rate of increase (Reynolds, 1997; proposed this reasoning to explain delays in the timing of the spring bloom associated with eutrophication). Positive anomalies in the meridional wind stress component (τ_v) were associated with delayed and stronger spring blooms along the European shelf (Fig. S20f). Changes in both spring and autumn/winter blooms in the transition zone were mainly associated with changes in SST and PAR, with a less important association with easterly winds and an important contribution of eddy kinetic energy along the Azores front (Fig. S20a–c). In contrast, wind stress (τ) was the main factor related to changes in the timing and intensity of autumn/winter blooms in the region presenting seasonal cycles with a single peak, especially in the Subtropical gyre (Fig. 5).

Discussion

We developed a novel approach to study changes in the seasonality of remotely sensed phytoplankton based on a probabilistic characterization of bloom

incidence. This allowed us to detect changes in the prevalence of different seasonal cycles and to propagate uncertainty in bloom detection to estimates derived from bloom statistics. We found a greater incidence of seasonal cycles typical of subtropical latitudes in the temperate North Atlantic, as well as changes in the timing and magnitude of blooms in the whole basin. Interannual variability in phytoplankton seasonality responded to environmental factors, especially to changes in wind patterns in lower and mid-latitudes, and sea surface temperature and incident photosynthetically active radiation in the northern North Atlantic.

Limitations and advantages of the methods employed to characterize seasonal changes in chl a concentration

A variety of approaches have been proposed to characterize phytoplankton seasonality using remote sensing data (Ueyama & Monger, 2005; Rolinski *et al.*, 2007; Platt & Sathyendranath, 2008; Thomalla *et al.*, 2011; Zhai *et al.*, 2011; Racault *et al.*, 2012; Sapiano *et al.*,

2012). Daily observations are usually aggregated and smoothed spatially and/or in time before analysis in an attempt to surpass the limitations imposed by data scarcity during cloudy periods (Gregg & Casey, 2007), which might otherwise result in huge biases in bloom statistics (Cole *et al.*, 2012 report typical errors of 30 and 15 days for the timing of onset and peak date in subpolar regions). Here we have preserved the daily time scale of chl *a* time series to minimize errors in bloom timing statistics, and instead tried to avoid problems derived from data gaps by increasing spatial coverage and, especially, by fitting a model to smooth available observations. We have not conducted a proper assessment of the effect of data gaps (e.g., Gregg & Casey, 2007; Cole *et al.*, 2012), but the relatively good agreement between estimates based on SeaWiFS and MODIS data indicates that this effect might remain low (the same cannot be ensured with respect to different chl *a* algorithms), especially considering the larger number of valid chl *a* retrievals provided by MODIS sensor each season. The detection of trends with different signs at the same latitudes also point in this direction (e.g., systematic biases usually consist in later bloom detections at high latitudes, see Cole *et al.*, 2012).

Analyses of phytoplankton phenology usually proceed by determining the occurrence or not of a bloom to later retrieve the date of onset, bloom magnitude and other statistics like bloom duration or mean chl *a* (e.g., Racault *et al.*, 2012). Frequently, the timing of the bloom is determined as the date when either modelled or observed chl *a* concentration reach the annual maximum or a threshold level which, in many cases, is determined based on a small fraction above median chl *a* concentration (e.g., 0.05). Both approaches assume that a bloom has occurred and thus, that chl *a* data contain enough information to characterize it. Here, we fitted a model with enough flexibility to characterize different types of seasonal cycles of chl *a* concentration (see Vargas *et al.*, 2008; Sapiano *et al.*, 2012). Seasonal chl *a* data were previously subset based on sea surface temperature time series. This allowed us to accommodate interannual changes in the shape of the seasonal cycle (Jönsson & Eklundh, 2002), and to determine directly the nature of blooms (i.e. spring *vs.* autumn/winter). An alternative approach consists in requiring chl *a* series to be above the threshold during two consecutive observations, a criterion which might be combined with setting a minimum variation in chl *a* to consider that a bloom might have occurred in a given location (e.g., Cole *et al.*, 2012).

The main difference with previous approaches consisted in avoiding the assumption that a bloom must have occurred and that it should be detected every year, i.e. ignoring changes in data availability or in the

type of seasonal cycle. Our approach was similar in this aspect to Sapiano *et al.* (2012), although it does not require a nearly constant seasonal cycle year after year at the same location to determine the lack or not of a seasonal cycle (see Vantrepotte & Mélin, 2009 for an alternative approach to the analyses of changes in chl *a*). Instead, we explored each oscillation in posterior simulations of models fitting available seasonal data. Observations retrieved during the target season – either spring or autumn/winter – were employed to estimate a threshold chl *a* concentration. The number of posterior simulations exceeding this selected chl *a* level was then used as an estimator of the probability of a bloom occurring or not. It should be noted that this procedure allowed us to detect changes in the prevalence of different seasonal cycles, to characterize uncertainty in bloom detection and, at the same time, to propagate this uncertainty to estimates derived from bloom statistics.

Changes in phytoplankton seasonality

We analysed changes in phytoplankton seasonality considering changes in the prevalence of different types of seasonal cycle, before examining changes in bloom timing and magnitude. In both cases, the length of the series poses a great limitation to ascribe observed trends to climate change (Henson *et al.*, 2010), although it revealed a clear signature of climate forcing on interannual changes in bloom statistics. Other problems include the difficulties to interpret changes in remotely sensed chl *a* concentration (see Materials and Methods), the indirect treatment of mixed layer dynamics and the lack of some important drivers of phytoplankton and bloom dynamics, like advection and sub-mesoscale features (Lehahn *et al.*, 2007; Mahadevan *et al.*, 2012). Another interesting aspect revealed by this study was the importance of taking into account potential differences between satellite missions and between chl *a* algorithms (see the Supporting Text in the Supporting Information).

In agreement with previous studies highlighting an increased prevalence of oligotrophic conditions (McClain *et al.*, 2004b; Polovina *et al.*, 2008; Irwin & Oliver, 2009), we found an increased prevalence of mean seasonal cycles with two peaks or even with a single autumn/winter bloom in the transition zone between regions presenting seasonal cycles with a single bloom, characteristic of subpolar and subtropical conditions. In the temperate North Atlantic, trends towards less apparent spring blooms contrasted with trends towards autumn/winter blooms of increased magnitude. At the same time, diverging trends in the timing of different peaks suggested an increase in seasonal stratification, the main consequence of increased surface warming

(Sarmiento *et al.*, 1998). Changes in bloom timing and magnitude might have also impacted several fisheries around the Atlantic (Platt *et al.*, 2003; Koeller *et al.*, 2009).

In the temperate North Atlantic, the association between changes in bloom metrics and environmental variables suggests that trends in bloom timing and magnitude reflect reduced light availability during winter and especially, an increased stratification due to surface warming and reduced wind stress (Henson *et al.*, 2009). Indeed, opposite trends in mean chl *a* on both sides of the basin in middle latitudes might be explained by different trends in wind stress. Weaker westerlies during winter might result in a decrease in mixed layer ventilation and nutrient renewal in the Northeast Atlantic, resulting in spring blooms of a reduced magnitude and a decrease in annual mean chl *a* concentration. These changes in seasonal wind patterns might be related to the widening of tropical wind circulation systems (Seidel *et al.*, 2007), and thus its effects on phytoplankton seasonality might continue and even strengthen in the next few decades.

Delayed and more intense blooms were found in most of the Subpolar gyre, although advanced and less intense blooms were more common in polar latitudes (Kahru *et al.*, 2011; Racault *et al.*, 2012). These changes were mainly associated with surface warming and light availability. In the Subtropical gyre, the different response on both sides of the North Atlantic can be explained again by changes in wind stress, although in this case the mechanism involved might be related to an increase in pigment cell levels in response to deeper mixing (Siegel *et al.*, 2005). Advanced and less intense blooms in the eastern side contrasted with delayed blooms of increased magnitude in the western subtropics, coinciding with trends of the same sign in mean chl *a*. All these changes in the seasonality of North Atlantic phytoplankton highlight the tight link between climate forcing and the dynamics of ocean ecosystems, stressing the value of remote sensing data for the monitoring, assessment, and projection of future climate change impacts on ocean ecosystems.

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Supporting Information

Additional Supporting Information may be found in the online version of this article:

Data S1. Detection and characterization of blooms: SeaWiFS vs. Aqua MODIS.

Data S2. Effect of chl *a* algorithm on incidences and derived estimates.

Data S3. Retrieval of chl *a* data and estimation of bloom metrics (extra discussion).

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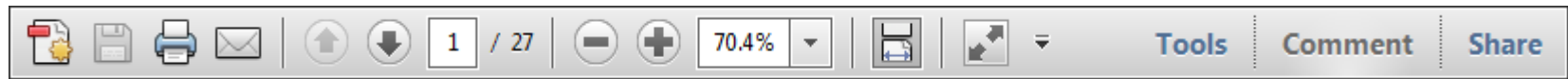
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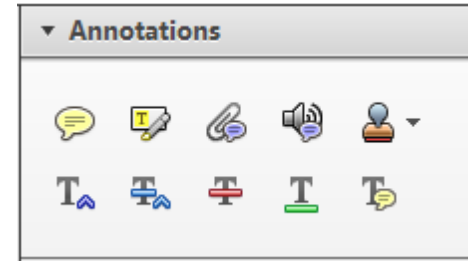
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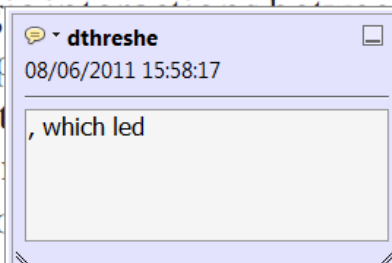


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standard framework for the analysis of microeconomics. Nevertheless, it also led to the emergence of strategic behavior in the number of competitors in an industry. This is that the structure of an industry, its main components, and the level, are exogenous to the industry. Important works on this by Shirasaka (henceforth) we open the 'black b



2. Strikethrough (Del) Tool – for deleting text.



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How to use it

- Highlight a word or sentence.
- Click on the **Strikethrough (Del)** icon in the Annotations section.

there is no room for extra profits and the number of competitors are zero and the number of firms (net) values are not determined by Blanchard and ~~Kiyotaki~~ (1987), perfect competition in general equilibrium. The effects of aggregate demand and supply in the classical framework assuming monopoly are an exogenous number of firms

3. Add note to text Tool – for highlighting a section to be changed to bold or italic.



Highlights text in yellow and opens up a text box where comments can be entered.

How to use it

- Highlight the relevant section of text.
- Click on the **Add note to text** icon in the Annotations section.
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4. Add sticky note Tool – for making notes at specific points in the text.



Marks a point in the proof where a comment needs to be highlighted.

How to use it

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USING e-ANNOTATION TOOLS FOR ELECTRONIC PROOF CORRECTION

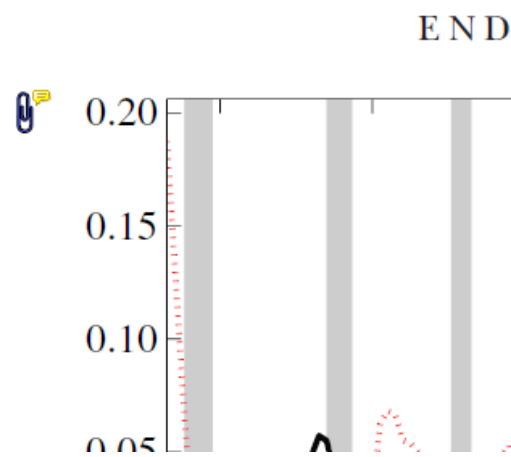
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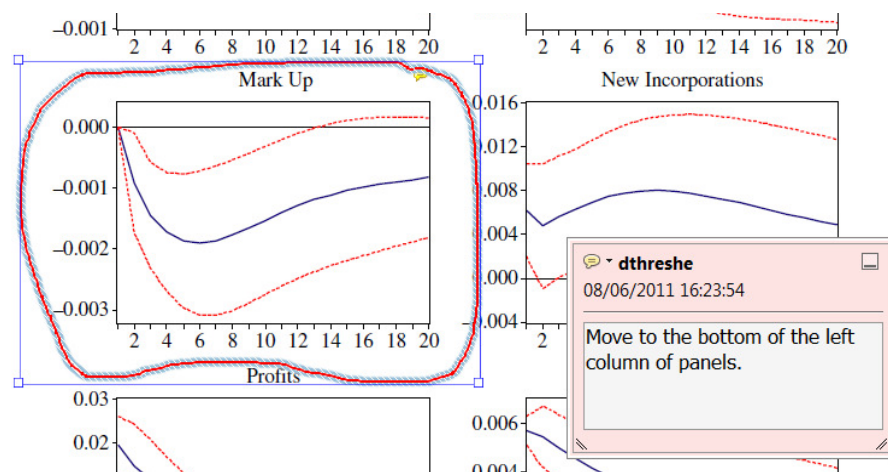


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